

# Did the Basel Process of Capital Regulation Enhance the Resiliency of European Banks?

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## Abstract

This paper analyses the evolution of the safety and soundness of the European banking sector during the various stages of the Basel process of capital regulation. We trace various measures of systemic risk and systematic risk as the Basel process unfolds and observe that, though systematic risk for European banks has been moderately decreasing over the last three decades, exposure to systemic risk has heightened considerably. This is particularly true, when we apply SRISK for the largest systemic banks. While the Basel process has succeeded in containing systemic risk for smaller banks, according to some measures it has been far less successful for the largest institutions. By exploiting the option of self-regulation embodied in the choice of internal models, the latter effectively seem to have increased their exposure to systemic risk as reflected in increasing SRISK. Hence, the sub-prime crisis found especially the largest and more systemic banks ill-prepared and lacking resiliency. This condition has even aggravated during the European sovereign crisis. Banking Union has not (yet) brought about a significant increase in the safety and soundness of the European banking system. Finally, low interest rates affect considerably to the contribution to systemic risk across the whole spectrum of banks.

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# 1 Introduction

The recommendations of regulating bank capital as developed by the Basel Committee of Banking Supervision were intended to increase the **safety and soundness** of banks and the global banking system while at the same time maintaining a **level playing field** in an increasingly globalized banking industry (Basel Committee of Banking Supervision, 1988).<sup>1</sup> Did these rules achieve their goals? How did they affect the conduct of banks in European banking markets where they were actually implemented?

After almost thirty years of capital reform, it is high time for an evidence-based evaluation of the economic consequences of the massive regulatory interventions since the initial adoption of the Basel Capital Accord in 1988 - now commonly referred to as “Basel I”. From an economic point of view this seems even more relevant, since the Basel Process was set into motion at a time of basically well functioning global banking markets, and without reference to any form of apparent market failure that required correction (see Goodhart, 2011)<sup>2</sup>. Subsequently, elements of self-regulation were introduced into the regulatory framework, initially to reward banks for implementing modern risk-management techniques in order to model market risks (market risk amendment of Basel I), and later even more prominently for modelling credit risk (“Basel II”). The evaluation of Basel II is complicated by the fact that the implementation period (2006) was immediately succeeded and, in fact, over-shadowed by the Great Financial Crisis and the subsequent repair operations now known as “Basel III” starting in November 2008.

This paper attempts a first evaluation of the Basel process of Capital Regulation. In particular, we track down the evolution of the resilience of European banks right from its implementation date and we identify its main drivers. By employing two recently developed measures of systemic risk, if resiliency was enhanced we would expect a long-run decline in the trajectories of these measures, reflecting an overall enhancement in the safety and soundness of European banking systems. We find differently.

The novel risk measures that we employ are currently widely discussed in academia and readily available in the public domain: i) SRISK has been developed by Brownlees and Engle (2017) and measures the capital shortfall (Acharya, Engle, Richardson, 2012), ii) Delta CoVaR developed by Adrian and Brunnermeier (2016) is built on comovements of asset prices. While SRISK can be viewed as an aggregate measures of banks’ *exposure* to systemic risk, Delta CoVaR is an aggregate measure of banks’ *contribution* to aggregate risk. Accordingly, SRISK is a measure of resiliency of a bank, and Delta CoVaR a measure of contagion risk originating in a particular bank.<sup>3</sup>

The SRISK measure is made publicly available for 100 European banks from 2000 onwards by Robert Engle’s V-Lab.<sup>4</sup> We reconstruct the measure back to the late 1980s, and for a larger sample

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<sup>1</sup>In its 1988 Report the Basel Committee on Banking Supervision explicitly states: “Two fundamental objectives lie at the heart of the Committee’s work on regulatory convergence. These are, firstly, that the new framework should serve to strengthen the soundness and stability of the international banking system; and secondly that the framework should be in [sic!] fair and have a high degree of consistency in its application to banks in different countries with a view to diminishing an existing source of competitive inequality among international banks.” (BIS, 1988)

<sup>2</sup>This is not saying that there was no issue of banking crises. Thus the implementation of Basel I coincides with the end of the S&L-crisis in the United States, which, however, according to general current perception is related to risk insensitive deposit insurance pricing. The argument that under-capitalized banks might be tempted to engage in risk taking has never been officially made by the Basel Committee.

<sup>3</sup>We present more precise formal definitions below in subsequent chapters.

<sup>4</sup><http://vlab.stern.nyu.edu/welcome/risk/>.

of listed European financial institutions, in order to trace its evolution all the way from the early days of the Basel Committee.<sup>5</sup>

Our central finding is that, in contrast to the ambitions of the Basel process, we cannot find a general reduction in the aggregate measures of systemic risk. Delta CoVaR is stationary across time, while SRISK is increasing in various stages that correlate with the introduction of internal models for market risk (Basel I, 1996) and again for credit risk (Basel II, 2006). Importantly we also identify heterogeneous effects across banking groups. The build-up in aggregate systemic risk SRISK is primarily driven by the highest quintile of large and systemic banks. If there are beneficial effects of capital regulation, they occur at the level of smaller banks in small economic magnitudes.

In retrospect, our results may not come as a big surprise. Academics have always been critical of the Basel process of Capital Regulation already at the time of inception of the various regulations. Hellwig (1995) was worried about correlations between market and credit risk not being properly addressed by Basel I regulation.<sup>6</sup> Danielson, Embrechts, Goodhart, Keating, Münnich, Renault and Shin (2001) raised serious concerns about the endogeneity of risks not being addressed at all within the Basel II framework, suggesting that Basel II might unintentionally and paradoxically even reduce safety and soundness of the banking system.<sup>7</sup>

But also on a purely methodological level the Basel approach attracted criticism of not accounting properly for extreme events and tail risk in particular. Notably, Eberlein and Keller (1995) demonstrate that hyperbolic Levy processes track real world market data far better than Gaussian processes. Building on this insight Eberlein et al. (1998) determine value-at-risk estimates and demonstrate that they tend to be much larger than under normality assumptions. At the standard 99%-VAR, typically Levy models would require double the amount of capital than Gaussian models would impose.<sup>8</sup>

Our main contribution is the identification of the main drivers of systemic risk in the cross-section of (European) financial institutions, and to pin down where Basel regulation has contributed to building up risks for the largest banks. To achieve this, we run panel and quantile regressions that allow us to control for macroeconomic factors as well as bank-specific balance sheet items. We then control for the implementation of important amendments of the Basel process of capital regulation, at the aggregate and at the bank level.

We observe that the implementation of market as well as credit risk models along Basel II regulation have a strong non-linear impact on individual bank systemic risk exposures as measured by SRISK. In quantile regressions, we find that the negative effect of internal market risk models is strongest for the largest and most risky banks, while it is small and less robust for the lower quantile. Furthermore, in a counterfactual exercise, we see that the historical observed level of SRISK was

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<sup>5</sup>In this sense our analysis also extends Engle, Jondeau and Rockinger (2014) who analyse 196 European financial institutions from 2000-2012.

<sup>6</sup>On a Panel Discussion on Capital Requirements for Market Risks Based on Inhouse Models in 1995 Hellwig (1996) suggests “that ten years later there may well be another panel, this one devoted to problems of quality assessment for inhouse models of credit risk and that a key question is what will happen to banks and banking systems in the ten intervening years”. History has replaced that panel with a true field experiment in the 2007-8 crisis. So this paper can also be viewed as a response to Hellwig’s (1996) request for an evidence-based evaluation of the internal model based approach to market risk.

<sup>7</sup>On the problem of neglecting the endogeneity of systemic risk see also Hellwig (2010).

<sup>8</sup>Incidentally, Eberlein et al. (1998) determine value-at-risk based capital for Deutsche Bank at more than double the amount required under the normality assumption of the market risk amendment (see their Table VIII).

in the last decade far higher than the systemic risk we would have had given the parameters in a pre-Basel II environment.

To investigate the motivations to such evidence, we study more attentively the effect of credit risk internal models with bank-level implementation data, and we observe a strong aggravating impact on SRISK from the implementation of advanced internal models. The results is particularly robust in a difference-in-differences approach carried out on the sample of banks implementing advanced IRBA models after Basel II, versus a sample of control banks matched by propensity score matching.

Overall we verify that the ECB, as a supervisor, was successful in identifying the most systemic banks when it took over supervisory responsibilities within the newly established European Banking Union in 2014. However, the systemic risk of those banks did not decline significantly after various stages of re-capitalisation and after entry into the Banking Union. Disconcertingly, both the recent innovations of Basel III as well as the regulatory attempts within the European Banking Union did not result in a significant decline in the aggregate SRISK measure. The individual SRISK of most systemic European institutions remains at levels considerably elevated relative to levels prior to the Great Financial Crisis of 2007-8.

With hindsight, we know that European banks were not well prepared for the Great Financial Crisis 2007-8. Capital buffers of large European banks were threatened to be wiped out even before the collapse of Lehman Brother. In the case of UBS, they proved to be as low as 1.4% (market) capital to asset ratios, while Deutsche Bank survived barely with a similarly low level of capitalization. Also in larger systematic analyses, Demirguc-Kunt, Detragiache, Merrouche (2013) and Beltratti, Stulz (2012) find that better capitalized banks were more resilient and fared much better during that crisis. Accordingly, it has been claimed (e.g. Gehrig, 2013a and 2013b) that the lack of capital, and hence resiliency, did result as an unintended consequence of the political process that did subsidize “optimizing” capital charges for individual banks and, ironically, so-called systemic banks in particular.

While the academic literature has focused on methodology and on developing new systemic risk measures, few studies exist about the market reactions to Basel-driven regulation in the banking industry. One important study is Wagster (1996) who has linked the Basel process to the competitiveness of banks across countries. In particular, he identified market reactions at various stages of the discussion about Basel I reforms. He shows that the Basel process can be viewed as a bargaining process between national regulators; many agreements by Japanese authorities, in particular concerning the regulatory treatment of hidden reserves, were elicited by concessions to the Japanese banking sector that were capitalized in market prices and can be measured accordingly. To the best of our knowledge our work is the first systematic evaluation of the effect of the Basel process of capital regulation on the safety and soundness of banking systems.

Moreover, concerning internal credit risk models we confirm and extend the results of Behn, Haselmann and Vig (2014) about the limitations of model-based regulation. Also they find negative effects of internal models for credit risk on the resiliency of German banks. We verify negative effects also for European banks and additionally we establish an important non-linearity of this effect across the SRISK-spectrum of banks.

Finally, we identify a considerable built-up of systemic risk in the insurance sector. We can trace this evolution back beyond the time span of other studies (Berdin, Sottocornola (2015), IMF Global Stability Report, 2016), starting in 1996 with the Market Risk Amendment of the Basel Accord,

and increasing in size and relevance thereafter. This findings are consistent with the existence of spillover effects from the banking sector to the insurance, following the change in regulation in the banking activity (Gehrig, Iannino, 2016).

The paper will proceed as follow: Section 2 briefly describes the Basel process. Section 3 introduces the data and the methodology used in the empirical analysis. Section 4 presents the main descriptive results and the main multivariate results. Section 5 highlights the policy role of our market-based measures. Finally, Section 6 concludes summarizing the unintended consequences of the Basel process.

## 2 The Basel Process

The Basel process of capital regulation was initiated in late 1974. The first meeting of Basel Committee on Banking Regulations and Supervisory Practices took place in February 1975. After a long period of consultations<sup>9</sup>, the first Basel Capital Accord (Basel I) was approved by the G10 governors in December 1987 and publicly announced in July 1988. The Accord was formally implemented in December 1992.

The Accord had already been amended in 1991, to reform the treatment of loan loss reserves, and later repeatedly in 1995 and 1996. The most important amendment was the introduction of internal models under supervisory review as an alternative to statutory rules in January 1996 as part of the market risk amendment (Basel Committee of Banking Supervision, 1996). This amendment essentially provided a choice between a self-regulatory regime under supervisory review and statutory regulation. It provided incentives to improve in-house risk management models, which were highly deficient in the 1990s even in multinational banks (see Wuffli, 1995). However, the amendment also implicitly provided incentives to employ internal models as an instrument to reduce regulatory burdens and capital charges, and, hence, to reduce resiliency (see Hellwig, 1995).

Proposals for a new capital accord were triggered by the initiation of a consultation process on a Revised Capital Framework in June 1999. This became the basis of the three-pillar framework of Basel II, which formally culminated in June 2006 in the agreement on Basel II: "International convergence of capital measurement and capital standards: a revised framework for comprehensive supervision".

Basel II was adopted in most countries with the notable exception of the U.S. However, the impact of its implementation could not be properly assessed<sup>10</sup> since already in 2007 the subprime crises developed into a worldwide crisis and depression. Hence, already in September 2008, the Basel Committee was forced to reconsider its regulatory framework with its guidelines on Principles for Sound Liquidity Risk Management and Supervision triggering the discussion on reforming Basel II, a process now commonly referred to as Basel III.

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<sup>9</sup>See Goodhart (2011) for details on the early years of the Basel Committee.

<sup>10</sup>Given the length of the consultancy process for Basel II, it is quite likely that the process did affect bank business models already well before the official implementation. Moreover, the self-regulatory pillar allowing internal models was available to officially and fully compliantly drive bank business models since 1996.

### 3 Sample and Methodology

In order to assess the implementation of the Basel principles, we propose an empirical investigation on a sample of European financial institutions from 1987 to 2015. The sample includes the listed institutions covered by Compustat Global and belonging to sector groups 4010 (banks), 4020 (diversified institutions), 4030 (insurance companies) and 4040 (real estate companies). To reduce survivorship bias, we include active as well as non-active institutions. We estimate systemic risk for 450 institutions from the Euro-area, Switzerland, and the United Kingdom, with at least 10 years of balance sheet data.

Compustat Global provides us with both daily market prices and capitalization and quarterly/annual fundamental data, such as book values of equity, assets and debt. As our quarterly data on European institutions go back till 1996, we complete the information back to 1987 with annual balance sheet data. We use the MSCI Europe index as the broad market return (Datastream data), and the yield on German federal bonds (Bundesbank data) as the risk free rate. Moreover, we use the market stress indicator CISS from Hollo, Kremer and Lo Duca (2012), and information on credit risk internal models from SNL Financial data, from the Bundesbank and from the Österreichische Nationalbank.

The empirical analysis proceeds as follows. We estimate the bank's exposure to systemic risk according to the SRISK measure proposed by Brownlees and Engle (2016) and its contribution to the aggregate systemic risk as the Delta CoVaR developed by Adrian and Brunnermeier (2016).

We analyse the impact of the Basel regulation on our measures of systemic risk considering both bank-level and market-level dummies for the use of internal model for credit risk (IRBA hereafter). SNL provides quarterly information on the implementation of standardized versus internal credit risk models from 2006, and we complete this information with the approval dates of internal models for all the German banks and the Austrian banks. Moreover, we use milestone dates that have affected the business models of the institutions in our sample: the introduction of market risk internal models in January 1996, the implementation of Basel II as June 2006, and the first publication of Basel III guidelines in September 2008.

We then control for a large set of potential drivers of the SRISK measure using information on either quarterly accounting bank data or weekly market data. We use bank cost of equity, market beta, market value, market-to-book ratio, total assets, investments in fixed income, investments in equity securities, non-performing loans, Z-score of distance to default. As market-level characteristics, we consider the MSCI Europe index, the short-term country policy rates, and the financial market stress indicator CISS.

Finally, we perform a counterfactual analysis to estimate the predicted SRISK with parameters pre-Basel environment, and a difference-in-differences analysis to pin down the effect of credit risk internal models on systemic risk.

#### 3.1 SRISK and Delta CoVaR Measures

We use measures of systemic risk that capt the contribution to aggregate risk of an institution in distress (Delta CoVaR) and the exposure in a distress market (SRISK and MES).

Building on the theoretical work of Acharya, Pederson, Philippon and Richardson (2017), the SRISK measure developed by Brownlees and Engle (2017) exploits market information to value expected capital shortfall, and, hence, the cost of recapitalizing a bank at any point in time. SRISK is a good measure of the exposure of an individual bank to systemic risk since it attempts to measure the (temporary) amount of capital a supervisor would have to inject to keep the bank running in an orderly way in case of a major disturbance. This systemic risk measure considers the combined effect of three components: sensitivity of the bank returns to aggregate shocks, leverage and size of the bank, and weakness of the financial system as a whole. Given these ingredients, a firm is considered systemically risky if it is likely to face a proportionally relevant capital shortfall when the financial sector is weak.

Adapting this methodology, we estimate (i) an asymmetric GJR GARCH model (Glosten, Jagannathan and Runkle, 1993) of the returns volatility of each institution and of the market equity index, (ii) a DCC correlation model (Engel, 2002) for the correlation between the institution return and the European market index, and (iii) the performance of the bank, and the capital shortfall in case of extreme financial downturns.

We assume a bivariate daily time series model of the equity returns of institution  $i$ , dependent on a value-weighted market index  $m$ :

$$\begin{aligned} r_{m,t} &= \sigma_{m,t}\varepsilon_{m,t} \\ r_{i,t} &= \sigma_{i,t}(\rho_{i,t}\varepsilon_{m,t} + \sqrt{1 - \rho_{i,t}^2}\xi_{i,t}) \end{aligned}$$

where the volatilities are asymmetric GJR GARCH processes and correlations are Dynamic Conditional Correlations (DCC). We use the MSCI Europe index for the market returns as a more representative benchmark for our sample of European banks.

The measures of performance and systemic risk are evaluated in the event of an extreme aggregate shock. We identify extreme downturns by falls in the daily market index higher than its 95% VaR. The expected daily loss of the bank returns, in this case, is the Marginal Expected Shortfall (MES):

$$MES_{it}(c) = E_{t-1}(r_{it}|r_{mt} < c = q_{5\%}) \quad (1)$$

The higher the bank's MES, the higher is its exposure to the risk of the financial system.

We estimate both the performance of the bank in such extreme events and the capital shortfall. The Long-Run Marginal Expected Shortfall (LRMES) is the expected loss in equity value of bank  $i$ , if the market were to fall by more than the above threshold within the next six months. As Acharya, Engel and Richardson (2012), we approximate it without simulation using the daily MES:<sup>11</sup>

$$LRMES = 1 - e^{(-18 * MES)} \quad (2)$$

Finally, given the above conditional expected equity losses, the current equity market value and the outstanding book value of debt, we determine the expected capital shortfall a bank would

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<sup>11</sup>The approximation represents the equity value loss over a six-month period conditional on a market fall by more than 40% within the next six months. Alternatively, we also estimate LRMES as NYU VLab currently updated to:  $1 - \exp(\log(1d) * \beta)$ , where  $d = 40$  is the six-month crisis threshold for the market index decline, and  $\beta$  is the dynamic market beta.

experience in case of distress. SRISK is defined as such capital shortfall in the event of an aggregate crisis:

$$\begin{aligned}
SRISK_{i,t} &= E_{t-1}[Capital\ shortfall_i|Crisis] \\
&= E_{t-1}[k(Debt + Equity) - Equity|Crisis] \\
&= E_{t-1}[k(Debt_{i,t}) + (1 - k)(1 - LRMES_{i,t})Equity_{i,t}|Crisis]
\end{aligned} \tag{3}$$

where:  $k$  is the prudential capital ratio, that we assume 8% in our main analysis. As robustness checks, we also conduct the inference analysis using capital ratio of 3% and 5.5%.

Once the individual  $SRISK_{i,t}$  are estimated per each bank, the relative exposure of firm  $i$  to the aggregate SRISK of the financial sector is:

$$SRISK\%_{i,t} = \frac{SRISK_{i,t}}{\sum_{j \in J} SRISK_{j,t}}, \text{ where } J = \text{firms with } SRISK > 0 \tag{4}$$

It represents the percentage aggregate capital shortfall that would be experienced by this firm in the event of a crisis, and it allows to identify the most systemic institutions in the sector.

Besides the estimation of the SRISK and its components, we also estimate the Delta CoVaR developed by Adrian and Brunnermeier (2016). It corresponds to the VaR of market returns conditional on a critical event on the returns of a bank  $i$ . The marginal contribution of bank  $i$  to the overall systemic risk,  $\Delta CoVaR$ , is the difference between the CoVaR in distress and the CoVaR in a median state.<sup>12</sup>

This measure starts from the estimation of an aggregate extreme loss in terms of Value-at-Risk, as the maximum loss of the market return within the  $\alpha\%$ -confidence interval, conditionally on some event  $C(r_{it})$  observed for bank  $i$ :

$$Pr(r_{mt} \leq CoVaR_t^{m|C(r_{it})}) = \alpha \tag{5}$$

Using a quantile regression approach, we identify this distress event of firm  $i$  as a loss equal to its  $(1 - \alpha)\%$  VaR:  $r_{it} = VaR_{it}(\alpha)$ .

The systemic risk of the bank  $i$  is then defined as the difference between the CoVaR of the financial system conditional on firm  $i$  being in distress and the CoVaR of the financial system conditional on firm  $i$  being in its median state:

$$\Delta CoVaR_{it}(\alpha) = CoVaR_t^{m|r_{it}=(VaR_{it}(\alpha))} - CoVaR_t^{m|r_{it}=Median(r_{it})} \tag{6}$$

Expressed in dollar terms, we weight it with the market capitalization of bank  $i$ :

$$\Delta^{\$} CoVaR_{it}(\alpha) = \Delta CoVaR_{it}(\alpha) * size_{it} \tag{7}$$

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<sup>12</sup>See Benoit et al. (2013) for a comparison of the two measures.



### 3.2 Drivers of Systemic Risk

Historically, prudential regulation of banks has focused on individual bank risk and micro-prudential regulation, neglecting the correlated systemic effects of several institutions in distress and in needs of recapitalization at the same time.<sup>13</sup> Therefore, it is of great interest to understand the potentially differential effects of the Basel measures on individual bank risk (microprudential risk) and aggregate systemic risk (macroprudential risk), and how the market prices systemic risk (Hellwig, 2009).

The analysis proceeds in the following steps. First, we perform a baseline regressions of SRISK with quarterly data on the overall sample, both as panel fixed effect and unconditional quantiles regressions. Next, we address the long-run relationship between average SRISK and market capitalization, distance to default and interest rate with a VECM on weekly averages. Thirdly, we introduce the analysis on the subsample of banks where microdata on the implementation on credit risk models is available. We use the full longitudinal information applying both panel fixed effects and quartile regressions to address non linearities in the relations between the systemic risk measure and the explanatory variables. Finally, we perform a counterfactual analysis to compare the observed SRISK on the systemic risk that would have been realized without Basel regulation. Many robustness checks are discussed separately in the Appendix.

In first baseline specifications on the full sample of 400 institutions, we regress the SRISK measure on Basel dummies, systematic risk proxies, market stress indicator, controlling for lagged quarterly firm characteristics  $Z_{kq-1}$ . In order to address issues of mixed frequencies, we aggregate the higher frequency measures (SRISK, Beta, market capitalization, CISS) to their quarterly median, and perform the following regression:

$$\begin{aligned} SRISK_{iq} = & \alpha + \gamma_0 SystematicRisk_{iq} + \gamma_1 CISS_q + \sum_k \gamma_k L.Z_{kqi} \\ & + \lambda_1 BIAmend + \lambda_2 BII + \lambda_3 BIII + \mu_i + \varepsilon_{iq} \end{aligned} \quad (8)$$

As measures of systematic risk, we consider alternatively (i) time-varying Beta between the bank asset returns and the market index, and (ii) cost of equity measured by means of a CAPM model.

The time-varying beta is estimated from the previous GJR-DCC GARCH model. The Cost of Equity is the CAPM return required by the market given the estimated dynamic beta and the risk premium of the market portfolio:

$$CostEquity_{it} = R_{ft} + \widehat{beta}_{it} * (R_{mt} - R_{ft}) \quad (9)$$

We use the daily annualized yield on German Bonds as a proxy for the risk-free rate for the European banks, and the MSCI Europe return compounded over the previous year as benchmark market return.

Concerning the Basel dummy variables, we perform our analysis considering three important milestones of the process in the time span of our sample: the introduction of Market Risk internal models (BIAmend, January 1996), the implementation of Basel II (BII, July 2006), and the guidelines of Basel III (BIII, September 2008). We therefore include three dummy variables identifying these subperiods: (i) Basel I and statutory regulation, (ii) Basel I and self-regulatory regime by use of internal market risk models, (iii) Basel II and self-regulatory regime concerning credit risk models, and (iv) Basel III.

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<sup>13</sup>Only under Basel III capital surcharges for systemically important financial institutions (SIFI) are introduced.

Finally, we control for market stress (CISS<sup>14</sup>), lagged total assets, non-performing loans, investments in equity security assets, investments in fixed income assets, market capitalization and leverage ( $Z_{kq-1}$ ).

Important non-linearities result from the graphical observation of the SRISK, therefore we regress the above specification both as a panel fixed effect model, and on three quantile regressions (at 0.25, 0.50 and 0.75 percentiles).

The second step of the analysis looks more carefully at the long-run relationship between SRISK, market capitalization, distance to default and policy interest rates. Distance of default is proxied by the Z-score (Boyd and Runkle, 1993, Fiordelisi and Marques-Ibanez, 2013). It measures the distance to becoming insolvent, as number of standard deviations away from the bank's ROA:

$$Zscore_{it} = \frac{ROA_{it} + E_{it}/TA_{it}}{\sigma_{ROA_i}} \quad (10)$$

We apply a VECM (with 4 lags) to weekly averages of the four variables, choosing the number of lags optimally given the AIC, HQIC and SBIC information criteria and assuring no autocorrelation is left in the residuals. To address breaks at the Basel introduction, we regress the VECM separately in three subperiods, before 1996, between 1996 and 2006, and from 2006 to 2015. Johansen tests of cointegration assure that we have one cointegrating equation in any period.<sup>15</sup>

Next, we focus on the third and more detailed part of the analysis. We introduce bank-level information on the credit risk internal models implementation (IRBA). This inclusion allows us to have a better understanding of how the usage of this regulatory tool has impacted on the systemic risk of the European banking sector. However, it reduces our sample size to 100 banks covered by SNL Financial and by the data provided by Bundesbank and Oesterreiche Nationalbank.

With the introduction of Basel II, banks were allowed to use in-house internal models to quantify risks of their loan portfolios instead of the standardized approach where the risk weights are assigned by coarse categories by the regulator. They have two options on how to implement IRBA, subject to authority approval: a foundation approach and an advanced approach. Under the former, banks are allowed to build their own models estimating the probability of default of individual clients or portfolios of loans. Under an advanced approach, banks can also estimate internally exposure-at-default and loss-given-default in order to quantify the risk-weighted assets.

We therefore regress weekly SRISK on a bank-level categorical variable,  $IRBA_{it}$ , equal to 1 for standardized models as Basel II regulation, 2 for the implementation of foundation internal models, 3 for mixed approaches, 4 for advanced IRBA models, and 0 before the introduction of Basel II regulations. With the introduction of this bank-level variables, we thus remove the time dummies identifying Basel II and Basel III. To account for the dynamic patterns we have evinced from the VECM, we regress a dynamic panel data model of SRISK, or we regress the residuals of a AR(2) model of SRISK on market- and bank-characteristics:

<sup>14</sup>We use the CISS measure as an indicator of the systemic stress that the European financial market as a whole experiences through time (Hollo, Kremer and Lo Duca, 2012). It consists in a weighted index of the instability in five different market segments: financial intermediaries, money markets, equities and bonds markets, and foreign exchange markets. The measure is publicly available from the ECB databank from 1999. We thank Manfred Kremer for sharing the CISS estimated from 1987.

<sup>15</sup>As a robustness check, we also proceed estimating possible lag-lead effects with a VAR between exposure to systemic risk (SRISK), contribution to systemic risk (Delta CoVaR), systematic risk (Beta), and market capitalization. We use 4 or 8 weeks to construct our VAR, and we test causality between these variables with Granger causality tests, however knowing we need caution about stationarity of SRISK. Results are reported in an online Appendix.

$$\begin{aligned}
SRISK_{it}^e &= \alpha + \sum_k \gamma_k L.Z_{kit} + \sum_q \gamma_q L.X_{qt} \\
&+ \lambda_1 BIAmend + \sum_p \lambda_p IRBA_{pit} + \mu_i + \varepsilon_{it}
\end{aligned} \tag{11}$$

We quantify the impact of proxies for systematic risk, using, alternatively, market Beta or CAPM cost of equity as previously estimated.

We control for market-characteristics ( $X_{qt}$ ) that would proxy for market investment opportunities (European market return, country policy rate, market stress indicator CISS), and bank-characteristics ( $Z_{kit}$ ) as market capitalization, intrinsic distance to default Z-score, and market over- or undervaluation of the bank (market-to-book ratio). Importantly, we apply country policy rates in order to control for monetary policy. We regress the panel model with fixed effects  $\mu_i$  for either bank or country.

Since we observe important nonlinearities in SRISK, we also use quantile regressions to address potentially differential effects of our covariates across the three main quantiles (q25, q50 and q75) of the distribution of the SRISK. We use an unconditional quantile approach as Firpo, Fortin and Lemieux (2009), where we marginalize the quantile coefficients using the recentered influence function. The interpretation of the estimated coefficients therefore corresponds to the usual interpretation as the marginal effect on the unconditional quantile of SRISK of a location shift in the distribution of the covariates, *ceteris paribus*.

### 3.3 Counterfactual analysis

The last step of the analysis is to estimate the counterfactual of systemic risk in the absence of Basel regulation. The complexity of this step is straightforward as all European banks are abiding to Basel regulation. We therefore provide two analyses to strengthen our previous findings: (i) we compare the impact of the market risk amendment and Basel II between different sub-sectors of financial institutions, banks versus insurers and real estates; (ii) we estimate the effect of Basel II in a difference-in-differences analysis where we observe treated banks with credit risk internal models versus control banks without treatment.

First, we apply the former weekly panel specification to estimate the parameters in two pre-Basel windows, separately for the three sub-groups of banks, insurance companies and real estates:

$$SRISK_{it}^e = \alpha + \sum_k \gamma_k L.Z_{kit} + \sum_q \gamma_q L.X_{qt} + \mu_i + \varepsilon_{it} \tag{12}$$

We have two estimation periods, as prior to the Market Risk Amendment (January 1996), and prior to Basel II regulation (June 2006). We do not re-estimate our parameters after 2008, because the effect of the crisis would confound our estimates.

Then, we predict what SRISK would have been in the following post-Basel periods, with the

assumption of constant parameters. We have two post-event windows where we predict SRISK as:

$$\widehat{SRISK}_{i\tau}^{(e,noMRA)} = \hat{\alpha} + \sum_k \hat{\gamma}_k^{(\tau < Jan1996)} L.Z_{kit} + \sum_q \hat{\gamma}_q^{(\tau < Jan1996)} L.X_{qt} \quad (13)$$

$$\widehat{SRISK}_{i\tau}^{(e,noBII)} = \hat{\alpha} + \sum_k \hat{\gamma}_k^{(Jan1996 \leq \tau < Jun2006)} L.Z_{kit} + \sum_q \hat{\gamma}_q^{(Jan1996 \leq \tau < Jun2006)} L.X_{qt} \quad (14)$$

Comparing the observed SRISK with the predicted  $\widehat{SRISK}$  for each financial sector helps us to reach a more robust interpretation of the effects of Basel regulation on the banks resiliency. In a well-functioning regulatory environment, we would expect to see a lower level for observed SRISK compared to the predicted SRISK only for the individuals affected, as such the banks and not other institutions.

Next, we use the variation in credit risk internal model implementation to discriminate between treated versus control banks, before and after Basel II implementation on June 2006. We implement a Difference-in-Differences analysis, such as:

$$SRISK_{it}^e = \sigma_0 + \sigma_1 IRBA + \sigma_2 BaselII + \sigma_3 BaselII * IRBA + \varepsilon_{it} \quad (15)$$

where IRBA is the dummy variable identifying banks with advanced or mixed credit risk internal models, BaselII is the time dummy capturing changes after the implementation of Basel II, and BaselII \* IRBA is the interaction term identifying IRBA banks after June 2006. The difference in differences parameter is therefore:

$$\hat{\sigma}_3 = (\overline{SRISK}_{IRBA,post} - \overline{SRISK}_{IRBA,pre}) - (\overline{SRISK}_{nonIRBA,post} - \overline{SRISK}_{nonIRBA,pre}) \quad (16)$$

In order to identify more precisely the control group and avoid selection bias, we perform first a kernel Propensity Score Matching (Rosenbaum et al. (1983)). We run a probit regression to estimate the probability of implementing IRBA models given market beta, Z-score, market capitalization, market-to-book ratio. A propensity score is then assign to balance the treated and the comparison groups. Next, the DD weighted regression is estimated, where observations are weighted to ensure that each group reflects the covariate distribution in the pre-Basel II period.

With this analysis, we remove both the bias in the post-Basel period between the treated and the control groups that could result from permanent differences between banks, and biases from comparisons over time in the treatment group that could be the result of other changes.

## 4 Results

We start our analysis by documenting the evolution of systematic as well as systemic risk. We will then illustrate cross-sectional properties of our systemic risk measures before analyzing the drivers of our measure of exposure to systemic risk (SRISK) by means of panel and quantile regression analyses. Next, we complement the analysis on SRISK by adding microdata on the implementation of internal credit risk models. In the fifth subsection, we report the results from the counterfactual analysis. In the subsection on robustness, we mention a variety of additional analyses not reported in detail in this paper. The underlying tables of that section are collected in an (Online-) Appendix. The last section concludes with a robustness analysis of the drivers of our measure of contribution to systemic risk Delta-CoVaR.

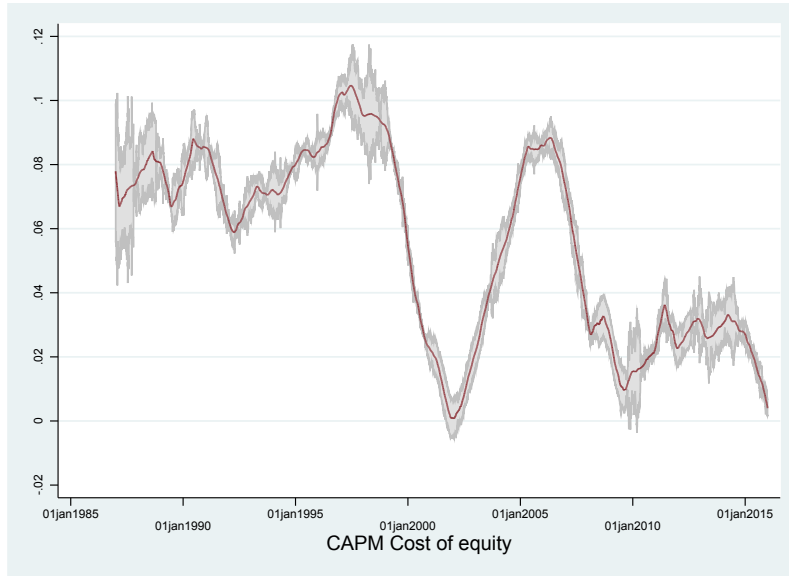


Figure 1: Evolution of CAPM cost of bank equity. The Figure presents the evolution of the daily average cost of equity across time, from January 1987 to 2015. We report a central moving average of 4 years, with confidence intervals around the estimated cost of equity. The cost of equity is the return required by the market applying a CAPM model with the time-varying beta and the annual risk premium required on the market return, as Equation 9:  $CostEquity_{it} = R_{ft} + \widehat{beta}_{it} * (R_{mt} - R_{ft})$ . We use the yield on German Bund as risk-free rate, and the MSCI Europe index compounded over the previous year as market return. The time-varying beta is estimated by a GJR DCC Garch model.

#### 4.1 Evolution of Systemic and Systematic Risk between 1987-2015

Arguably, the Basel process of capital regulation was intended to increase the stability and safety of the banking industry.<sup>16</sup> Accordingly, we might expect to see a decrease in the riskiness of bank business models after the formal implementation of the various stages of the Basel process. This should ideally be reflected in lower risk premia and, hence, lower funding costs such as lower costs of bank equity. Moreover, we would expect that an improvement in stability and safety of the banking sector would also be reflected in a reduction in the exposure of banks to systemic risk. So what do the data tell us about the evolution of these measures for (almost) the past three decades?

Indeed, it turns out that market-based measures of the cost of bank equity did decrease significantly on average (Figure 1).<sup>17</sup> The trend is pervasive across countries and we do not observe particularly different time trends (Figure 2). This decrease did happen across most of the advanced economies such as Germany and Great Britain, consistent with previous literature on the G-10 countries and the US in particular (Baker and Wurgler, 2015, Maccario et al. 2002).

While permanently declining long-term interest rates contributed to a decline in cost of bank equity, most of this decrease is actually driven by an average decline in systematic risk, as measured

<sup>16</sup>Gehrig (1995) suggests that harmonization of regulation and creating a level playing field was another goal of the Basel Committee.

<sup>17</sup>We follow the Federal Reserve System (Barnes and Lopez, 2006) and the BIS (King, 2009) approaches in measuring costs of bank equity on the basis of a CAPM-model. Moreover, again following King (2009) we provide a moving-average presentation of the daily cost of capital, which is notoriously volatile. Unlike King (2009) we provide confidence intervals to allow an assessment of statistical significance.

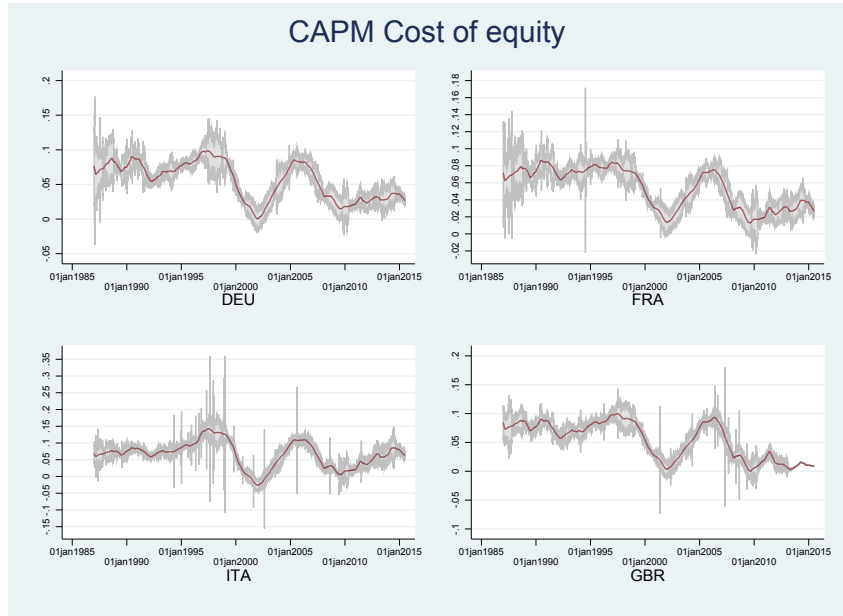


Figure 2: Cross-country variation of CAPM cost of bank equity. The Figure presents the evolution of the daily average cost of equity by different countries, across time, from January 1987 to 2015. We report a central moving average of 4 years, with confidence interval. We report Germany (DEU), France (FRA), Italy (ITA) and United Kingdom (GBR). The cost of equity is the return required by the market applying a CAPM model with the time-varying beta and the annual risk premium required on the market return, as Equation 9.

by the beta (Figure 3). We observe significant country variation (Figure 4), however, with the main exception of Italy, systematic risk appears as being in long-term decline on average. This finding implies increasingly more favorable funding conditions for banks and lower costs of issuing bank equity.<sup>18</sup> It is also worth noting that this long-run evidence runs counter to the perception of contemporary observers in the early phase of the Basel process. The ubiquitous sense of increasing risk in banking due to narrower intermediation margins caused by deregulation and intensified competition (e.g. Gehrig, 1995, 1996) is not reflected in average risk premia across European banks in subsequent decades.

How does this evidence of lower risk premia and systematic risk in the financial industry relates to measures of systemic risk? Has the safety and stability of the banking system been enhanced by the Basel process at large? Are there segments in the banking system where resilience has actually declined?

To address these questions, we employ two standard measures of systemic risk, the Delta CoVaR measure of Adrian and Brunnermeier (2016) as a measure of an institution's contribution to systemic risk, and the SRISK measure of Brownlees and Engle (2016) as a market-based measure of an institution's exposure to systemic risk.

The Delta CoVaR measure first peaks in the late 1980's at the end of the S&L crisis. After the Basel accord of 1988 the Delta CoVaR measure is in decline until 1996, from which on it remains

<sup>18</sup>This finding is in line with Barnes and Lopez (2006) and King (2009) who also find downward trending cost of bank capital except for the case of Japan.

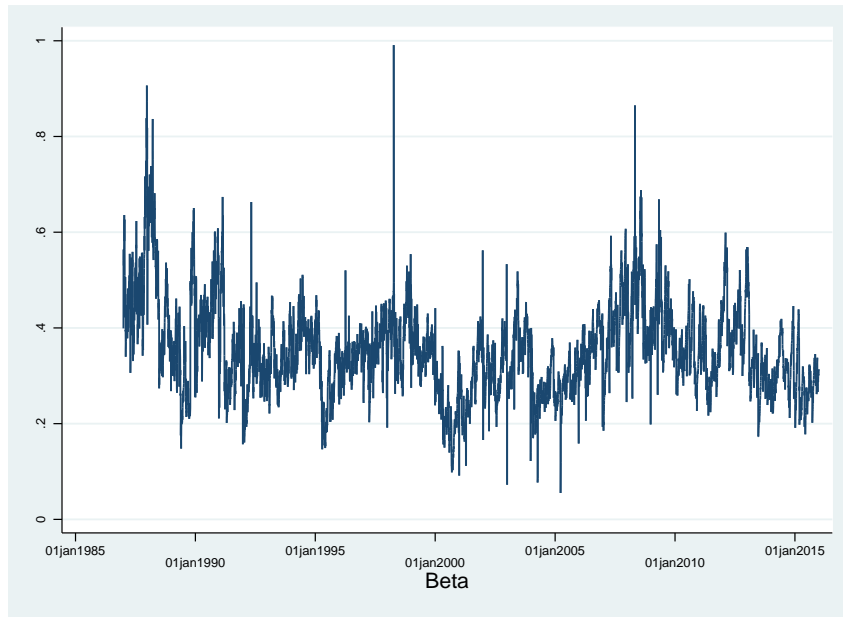


Figure 3: Evolution of systematic risk (Beta). The Figure presents the evolution of the daily average beta across time, from January 1987 to 2015. It represents the sensitivity of the bank equity returns to the MSCI Europe index returns. It is estimated by a GJR DCC Garch model.

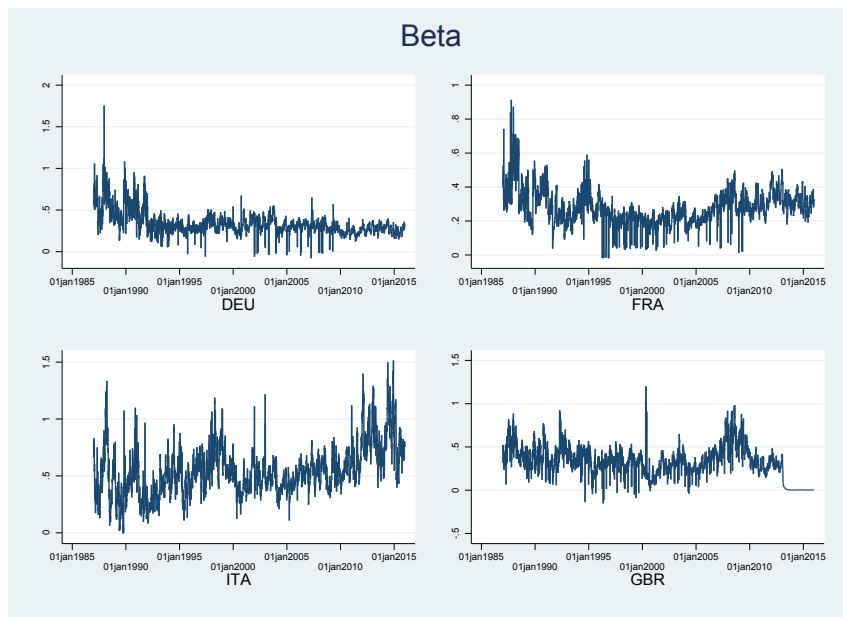


Figure 4: Cross-country variation of systematic risk (Beta). The Figure presents the evolution of the daily beta by different countries, across time, from January 1987 to 2015. We report Germany (DEU), France (FRA), Italy (ITA) and United Kingdom (GBR). It represents the sensitivity of the bank equity returns to the MSCI Europe index returns. It is estimated by a GJR DCC Garch model.

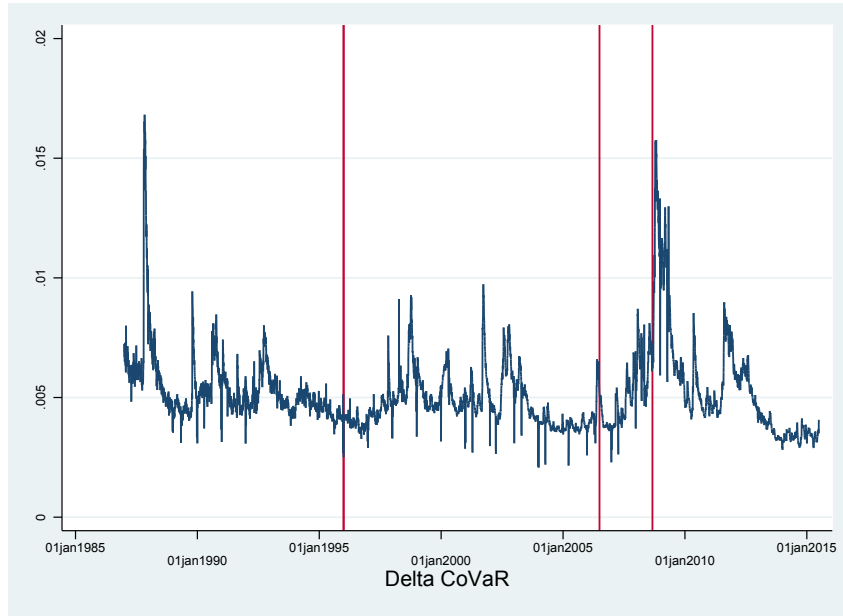


Figure 5: Evolution of contribution to systemic risk - Delta CoVaR. The Figure reports the evolution of the daily average estimated Delta CoVaR in Equation 6, estimated with quantile regressions.

heightened until about 2005 (Figure 5). This period roughly corresponds with the period after the introduction of the market risk amendment of Basel I until the end of the consultancy process of Basel II. This period also covers the dot-com bubble, which apparently did not affect contagion risk of European financial institutions. The next huge increase in Delta CoVaR coincides with the European sovereign crisis in 2009-10.

Surprisingly, while figuring significantly, the subprime crisis does not figure prominently according to the Delta CoVaR measure. There is a single peak around Lehman failure in September 2008, but Delta CoVaR remains below pre-crisis levels. To the effect that the subprime crisis has been characterized by a drying-up of liquidity, it appears remarkable that contagion risk has not shot up dramatically during the 2007-8 period.

The SRISK measure exhibits a markedly different pattern. In fact, we present two versions: i) in the first version we aggregate over shortfalls and surpluses of individual banks (Figure 6), and ii) in the second version we only aggregate positive shortfalls (Figure 7). While the first version does implicitly allow for inter-industry netting of bank capital, the second version measures the total amount of re-capitalization needed for a given capitalization standard. Thus, the net measure is a measure of the shortfall from a societal level after potential redistribution of bank capital, while the latter measure is an indicator of overall industry stress.

Overall, both measures exhibit an increasing trend suggesting a growing reduction in resilience. Both SRISK measures are quite low around the dot-com bubble, which may just be a reflection of the bubble per se.<sup>19</sup> The measure shoots up when the bubble bursts, but remains elevated prior to the subprime crisis. During the Great Recession it shoots up again after the Lehman failure, but

<sup>19</sup>Since SRISK is a market-based risk measure it underestimates true exposure to systemic risk in periods of



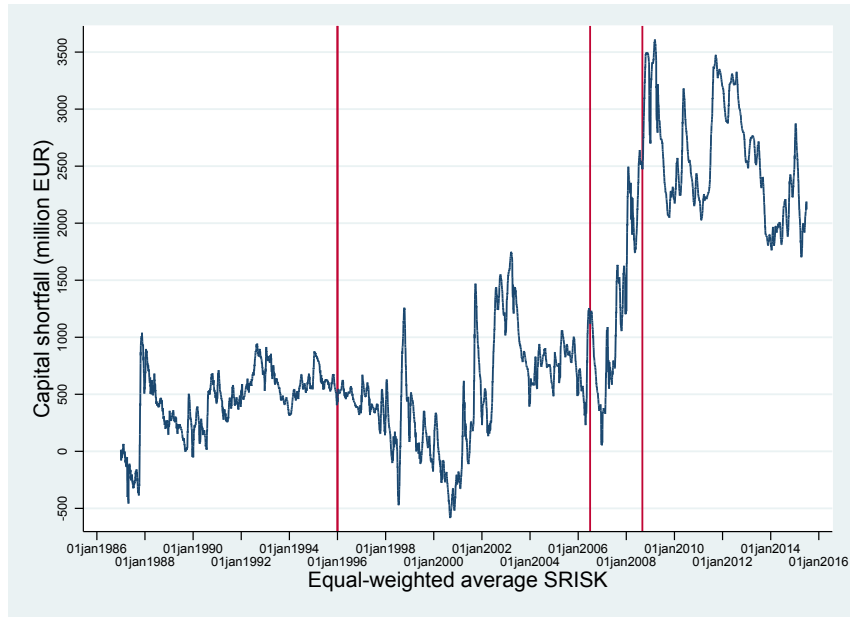


Figure 6: Evolution of exposure to systemic risk - average SRISK. The Figure reports the evolution of the daily average estimated SRISK in Equation 3. We report a central moving average of 20 days. We consider both positive and negative values of SRISK, respectively as shortfall and surplus of capital. The SRISK is estimated by MLE using a GJR-DCC Garch model. We use a capital ratio  $k=8\%$ .

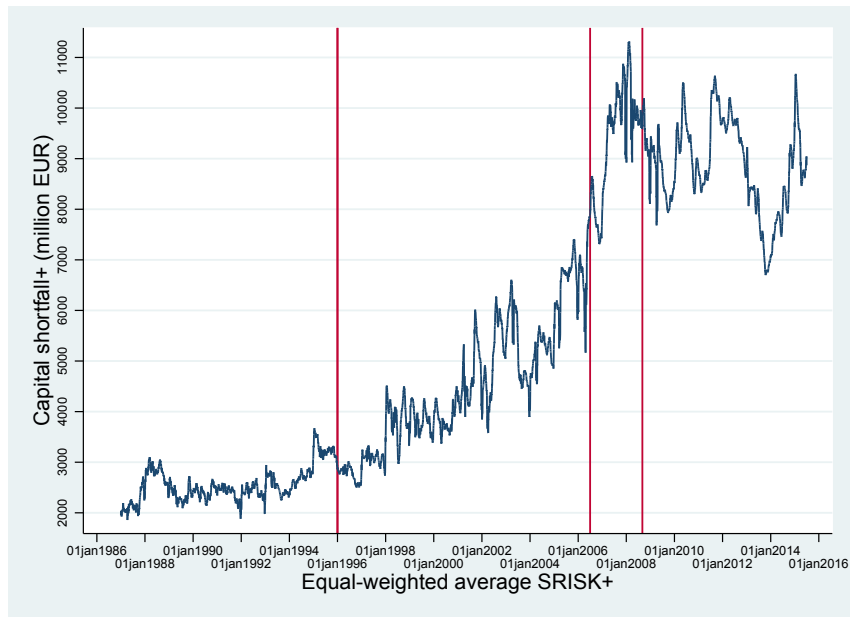


Figure 7: Evolution of exposure to systemic risk - average positive SRISK. The Figure reports the evolution of the daily average estimated SRISK (Equation 3) in case of capital need (positive SRISK). We report a central moving average of 20 days, and we consider only positive values of SRISK, representing the capital shortfall in the system. The SRISK is estimated by MLE using a GJR-DCC Garch model. We use a capital ratio  $k=8\%$ .

subsequently, and in contrast to Delta CoVaR, remains at almost identically high levels during the European sovereign crisis. On the positive side though, the Basel III measures seem to be effective in preventing a further rise in SRISK, albeit at a rather high level well above that observed in the 1990s.

By visual inspection of the first SRISK measure three major level changes in aggregate SRISK catch the eye: i) the early stage from 1988-2001, ii) the period from 2002-2009 and the iii) sovereign crisis stage from 2009 onwards. Again the liquidity crisis of 2007-8 does not exhibit dramatic effects on SRISK.

While SRISK clearly indicates a sharp build-up of exposure to systemic risk, also Delta CoVaR does not provide any evidence of a reduction in the contribution to systemic risk, and, hence, an increase in resiliency in the banking system. According to these measures, the Basel process does not seem to have achieved the goal of increasing the stability and safety of the banking system relative to the pre-Basel era, at least for the European countries.

## 4.2 Evolution across banks and sectors

These results lead to a fundamental question: Can we say anything about the sources of the build-up of systemic risk? We first check whether the build-up has been uniform across the banking system or whether it is related to particular institutions or business models.

When looking into quintiles of the SRISK measure it turns out that it is the upper two quintiles that massively build up SRISK, while in the case of CoVaR, banks seem to be more uniformly affected. Accordingly, important non-linearities show up in the case of SRISK (Figures 8 and 9).

It is interesting to note that the introduction of internal market risk models in 1996 did exert a short-lived but discernible moderating effect on the SRISK-trajectories across all quintiles (Figure 8).

Moreover, we also show that the banks that more contribute to SRISK are the banks now supervised directly by the ECB within the new Banking Union regime officially in place since November 2014 (Figure 10). We can see that SRISK has increased across all quintiles and also for banks not supervised directly by ECB, reconfirming an overall increase in systemic risk according to that measure. However, the banks chosen by the ECB are the ones that have distinguished themselves in building up a massive contribution to the systemic risk after 2008.

Let us now take a system's perspective on the whole financial system by differentiating according to banks (1), diversified financials (2), insurance companies (3) and real estate companies (4). Figure 11 reports the evolution of the number of institutions belonging to each group (according to the Compustat Global classification).

Figure 12 establishes that there is significant variation in the time trend of Beta for the different sectors. It is particularly the diversified financials and the insurance sector that saw a decline in beta, while the banking sector per se experiences a rise in beta after the financial crisis.

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overpricing (bubbles) and it overestimates true exposure to systemic risk in periods of underpricing. In this sense SRISK is not a useful early warning indicator.

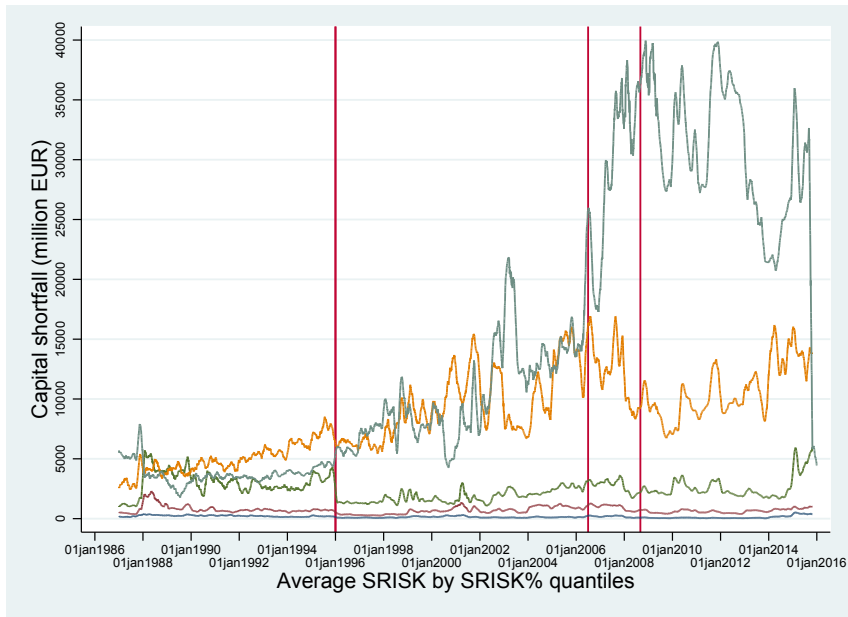


Figure 8: Quantile effects and non-linearities. The first frame reports the evolution of the daily average estimated SRISK (Equation 3), distinguishing five equal-size quintiles of contribution to capital shortfall (SRISK%), as in Equation 4. The top quintile (gr5) corresponds to the group of banks with the highest level of positive SRISK, while the bottom quintile (gr1) corresponds to the group of banks with the lowest level of capital shortfall. We report a central moving average of 50 days, and we average both positive and negative values of SRISK.

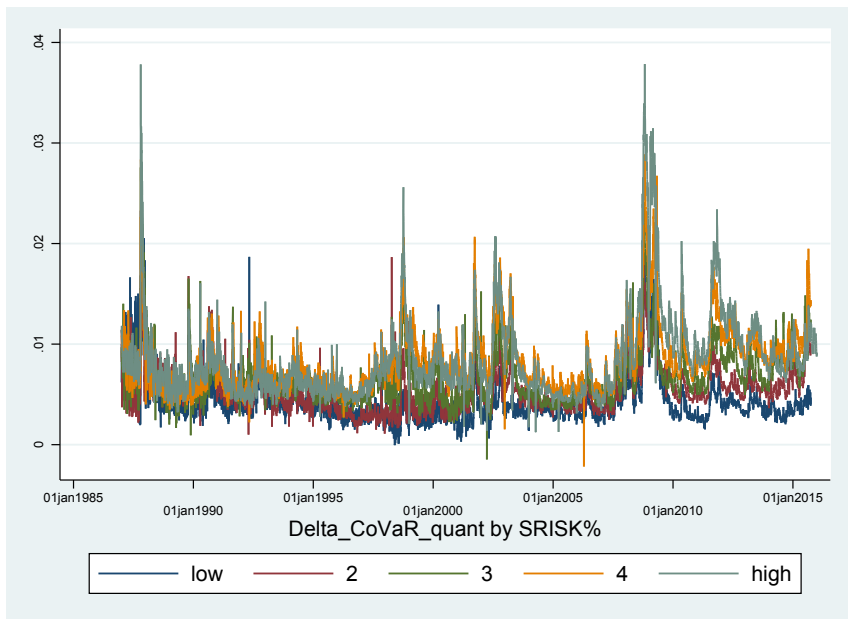


Figure 9: Quantile effects and non-linearities. The Figure reports the evolution of the daily estimated Delta CoVaR (Equation 6), distinguishing for the five quintiles of SRISK%, as in Equation 4. The top quintile (gr5) correspond to the group of banks with the highest level of CoVaR, while the bottom quintile (gr1) correspond to the group of banks with the lowest level.

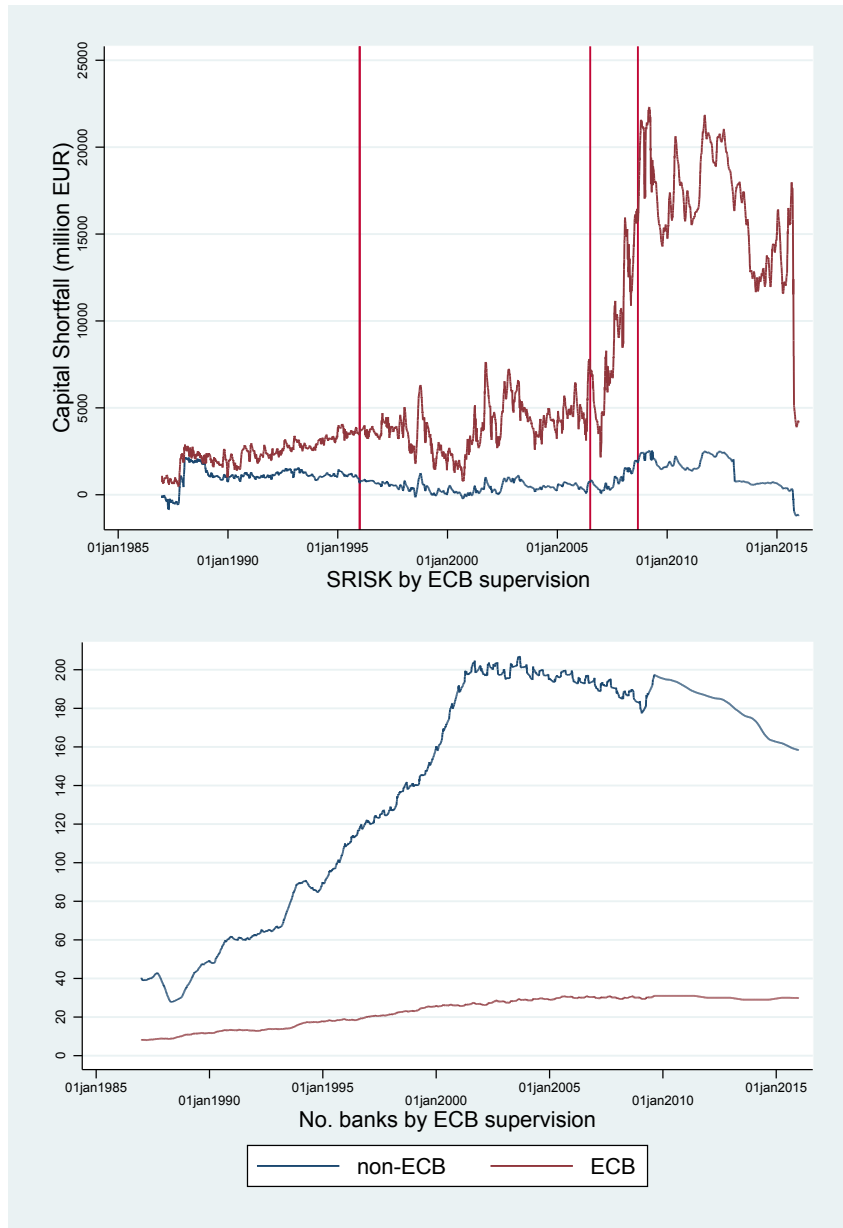


Figure 10: ECB Supervision. The first frame reports the evolution of the daily estimated SRISK (Equation 3), distinguishing for ECB supervision. The first group (ecb0) correspond to the group of banks that were not part of the Stress Test performed by ECB in 2014, while the second group (ecb1) correspond to the group of banks that were part of the ECB Stress Test in 2014 and their supervision has been taken over by ECB. The bottom frame reports the evolution of the number of banks in the the two groups, under ECB supervision from 2014 and not, followed backward to 1987. We report a central moving average of 1 year.

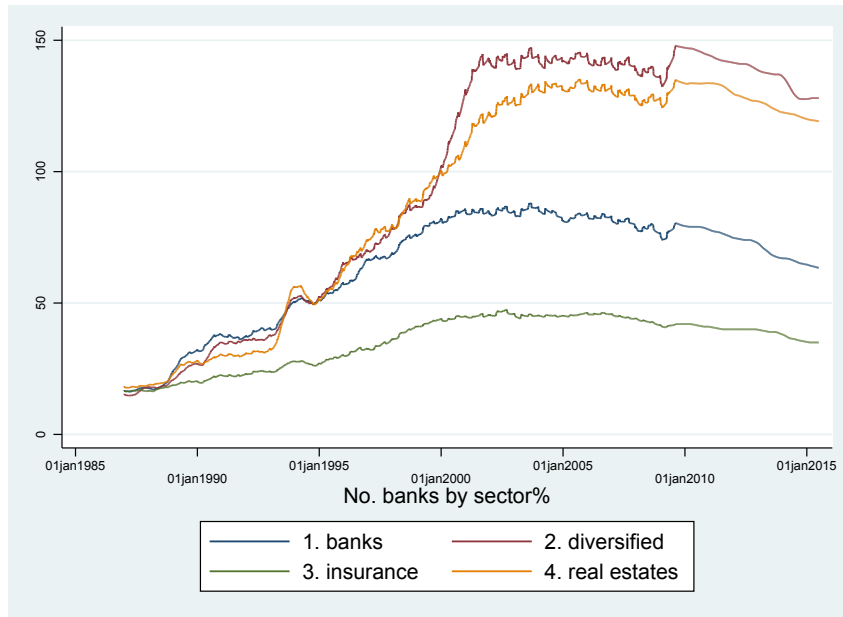


Figure 11: Number of banks. The Figure reports the evolution of the number of banks in the sample from 1987 to 2015, by financial sector. We report a central moving average of 1 year. We report banks (1), diversified financials (2), insurance companies (3) and real estates (4)

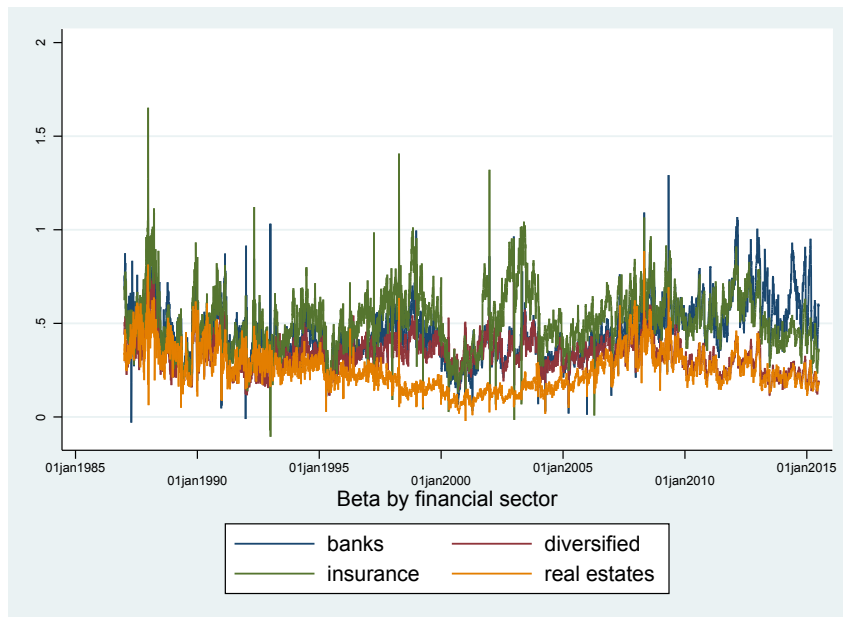


Figure 12: Evolution of measures of systematic risk. The Figure presents the evolution of the daily average beta across time, from January 1987 to 2015. It represents the sensitivity of the bank equity returns to the MSCI Europe index returns. It is estimated by a GJR DCC Garch model. We report banks (1), diversified financials (2), insurance companies (3) and real estates (4).

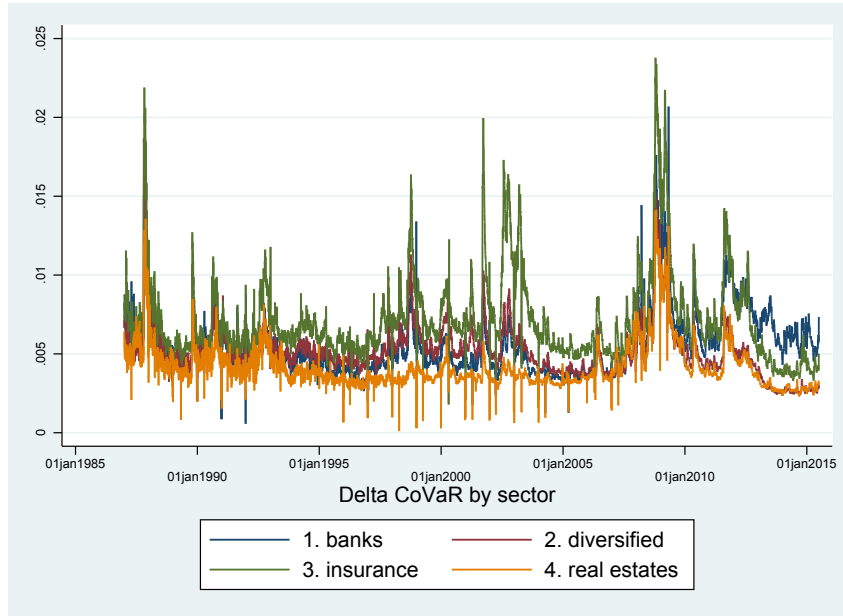


Figure 13: Evolution of systemic risk - Delta CoVaR. The Figure reports the evolution of the daily average estimated Delta CoVaR in Equation 6, by sector. We report banks (1), diversified financials (2), insurance companies (3) and real estates (4).

Interestingly Figure 13 demonstrates that the risk of contribution to systemic risk, i.e. contagion risk, has been largest in the insurance sector. Also Delta CoVaR for the insurance sector follows Delta CoVaR in the banking sector with a small lag. Moreover, Delta CoVaR has been rising between 2002-2005 and again in the European sovereign crisis in 2009-2010. The real estate sector exhibits significant contagion risk only during the European sovereign crisis. Overall, however, Delta CoVaR measures are quite generally highly correlated across sectors.

The most striking differences across intermediary groups can be witnessed in the exposure to systemic risk as demonstrated in Figure 14. SRISK has considerably increased both for banks, around 2001 and again after 2009, and for insurance companies especially after 2002. Moreover, exposure to systemic risk is still increasing in the case of the insurance sector.

To summarize: according to the SRISK measure, the overall exposure to systemic risk has increased considerably after the implementation of the Basel accord. This finding may seem surprising, since it suggests that the Basel process of capital regulation has failed to achieve its stated goals of increasing the safety and stability of the banking system at large. This conclusion is less strict by employing Delta CoVaR as a measure of contributing to systemic risk. However, also Delta CoVaR certainly does not suggest a general reduction of contagion risk during the various stages of the Basel process of capital regulation. The evolution of these systemic risk measures runs counter to the evolution of measures of systematic risk, which appears to be on a declining long-term trend for most European countries, albeit at very slow pace.

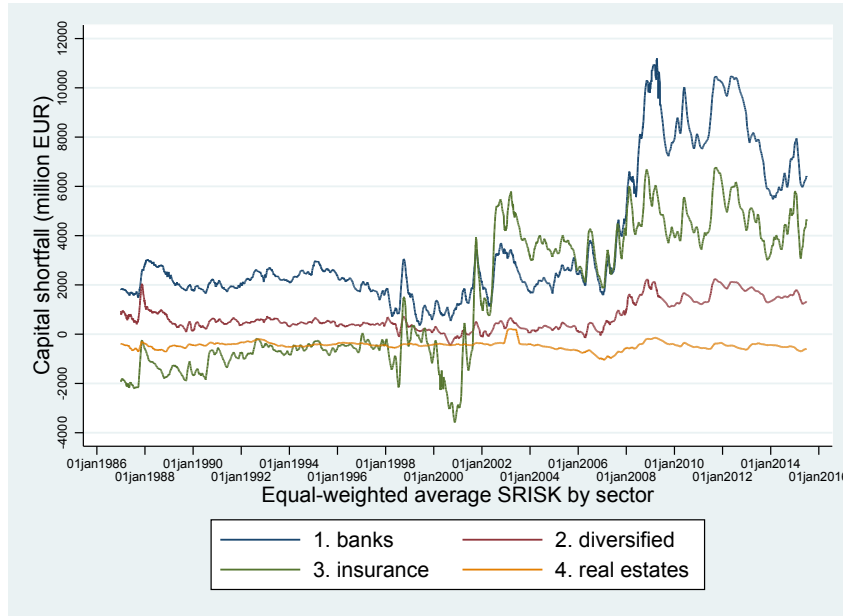


Figure 14: Evolution of systemic risk - SRISK. The Figure reports the evolution of the daily average estimated SRISK in Equation 3. We report a central moving average of 20 days. We consider both positive and negative values of SRISK, respectively shortfall and surplus of capital. The SRISK is estimated by MLE using a GJR-DCC Garch model. We report banks (1), diversified financials (2), insurance companies (3) and real estates (4).

### 4.3 Drivers of Systemic Risk

While the bivariate analysis reveals a systematic and permanent increase in the SRISK measure, it does provide only limited information about the potential drivers of these developments, such as size and other balance sheet items. To analyze the causes of the increase in systemic risk, a fully fledged multivariate analysis is required. We regress the estimated SRISK on different cost of equity measures, as well as lagged balance sheet items, including fixed effects and year dummies.

Table 1 shows the first motivating results. SRISK is positively affected by the composite indicator of systemic stress in the overall system and by the size of the total assets, in all specifications. Therefore, controlling for market stress and total assets, we observe that SRISK is contained, by construction, by the market capitalization of the bank. Non-performing assets are positively related, but unexpectedly not significant in mean. Systematic risk, as measured by the market beta, is also a significant driver of SRISK, while return on equity (ROE) does not play a significant impact on the systemic exposure. Counter-intuitively, the cost of bank equity, on the other hand, has a strongly moderating effect on systemic risk. Evidently, it is the low-cost banks that impose the largest contribution to SRISK. We have previously seen that the cost of equity has strongly declined in mean from 2000 to 2010, just when the exposure to systemic risk has heightened.

Looking at the Basel process, we can interestingly observe that the market risk internal models dummy remains insignificant in these specifications, suggesting that the market risk amendment of January 1996 has indeed not reduced systemic risk. Moreover, the Basel II dummy (July 2006) has contrasting results across specifications, showing only a weak contribution in reducing systemic risk and suggesting a positive and significant risk increasing role in the third specification, probably due

to the drop in cost of equity in 2000s. These reforms were steered towards improving capital and liquidity positions on banks' balance sheets. The Basel III dummy remains inconclusive.

Surprisingly, cost of equity has a moderating effect on SRISK. This contemporaneous effect could, however, result from endogeneity. Well capitalized banks will enjoy lower cost of bank equity and for that reason exhibit lower SRISK. We therefore focus most of the next analysis on the Beta as proxy for systematic risk.

Next, results on the policy variables could be driven by the non-linearities that we have observed earlier in the discussion of the SRISK measure.

In order to control for potential non-linearities in SRISK, we also present quantile regressions in Table 2, and indeed the quantile regressions contribute to resolving the puzzle on the mean estimations. It turns out that both regulatory instruments, the market risk amendment of Basel I (January 1996) and the implementation of Basel II (July 2006) significantly contributed to reduce systemic risk in the lower quartile, while they affect SRISK gradually less moving to the upper quartiles. The introduction of modelling market risk internally has effectively increased the contribution to systemic risk in the upper quartile.

While internal models for measuring market risk contributed to increase resiliency in the safer segment, they have contributed to a build-up of systemic risk precisely in the systemically relevant segment. Hence, after 20 years of evidence, it turns out that Hellwig's (1995) concerns proved right, while Wuffli's (1995) optimism was unfounded. With hindsight, we witness that it was the smaller banks that reduced their systemic risk contribution while the large banks (mis-)used the options offered by the use of internal models to effectively enlarge their systemic risk exposure and reduce their resiliency. Clearly this effect appears as an unintended consequence of the 1996 amendment. The Basel II dummy suggests a significant risk-reducing effect only in the lowest quartile. In the higher quartiles the moderating effect becomes economically smaller and statistically insignificant. These motivate the analysis in the next section, where we introduce detailed micro-econometric information about the implementation dates of internal models.



Table 1: Drivers of SRISK (quarterly panel regression)

SRISK (k=8%)	(1)	(2)	(3)
Beta	1,459*** (311.9)		
ROE		2.347 (2.645)	
Cost of Equity			-2,077** (909.5)
CISS	1,550*** (323.9)	1,717*** (370.9)	754.9*** (276.7)
Tot.Assets	0.0688*** (0.00849)	0.0684*** (0.00916)	0.0686*** (0.00865)
NPA	0.135 (0.135)	0.141 (0.147)	0.147 (0.142)
Equity Securities	-0.0297 (0.0289)	-0.0496 (0.0401)	-0.0317 (0.0304)
Fixed Income Securities	-0.00536 (0.00588)	-0.00683 (0.00890)	-0.00315 (0.00611)
Mark.Cap	-0.569*** (0.0750)	-0.537*** (0.0778)	-0.561*** (0.0788)
Leverage	0.646** (0.252)	-0.344 (0.227)	-0.356 (0.233)
Market Risk Amendment	-175.2 (505.1)	1,041 (654.4)	1,001 (610.1)
Basel 2	-39.07 (34.99)	48.12 (36.09)	80.34** (37.21)
Basel 3	43.04 (126.7)	-74.41 (142.8)	-130.1 (130.8)
Constant	-1,217** (483.6)	-944.9* (499.2)	-700.3* (392.3)
Year effects	yes	yes	yes
Firm fixed effects	yes	yes	yes
Observations	16,746	13,532	16,745
R-squared	0.825	0.824	0.823
Number of id	400	383	400

Clustered standard errors in parentheses

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1

<sup>a</sup> This table reports the results from the quarterly panel regressions of SRISK with fixed effects and year dummies for the individual banks. We regress the SRISK measure on CISS (indicator of systemic stress in the European system, Hollo et al, 2012), cost of equity measures, such as Beta (model 1), CAPM cost of equity (model 3) or ROE (model 2), an internal model dummy from January 1996 and two Basel dummies, Basel II from June 2006 to September 2008, and Basel III from September 2008. The Beta is estimated from a GJR-DCC Garch model between the bank stock returns and the MSCI Europe index. The CAPM cost of equity is the return required by the market applying the time-varying beta to the annual risk premium required on the market return, as Equation 9:  $CostEquity_{it} = R_{ft} + \widehat{beta}_{it} * (R_{mt} - R_{ft})$ . We use the yield on German Bund as risk-free rate. Moreover, we include drivers as total assets, non-performing assets, other investment in equity and fixed income securities, market capitalization and leverage.

Table 2: Quantiles Regressions (quarterly)

SRISK (k=8%)	(1) Q.25	(2) Q.50	(3) Q.75
Beta	108.6*** (13.66)	76.34*** (10.34)	36.62*** (10.80)
CISS	103.5*** (16.00)	86.89*** (12.29)	63.96*** (12.12)
Tot.Assets	0.0791*** (0.000728)	0.0822*** (0.000308)	0.0854*** (0.00289)
NPA	0.00915*** (0.00233)	-0.00651* (0.00368)	-0.0200 (0.0544)
Equity Securities	0.0247*** (0.00308)	0.0314*** (0.00148)	0.0313 (0.0388)
Fixed Income Securities	-0.00359*** (0.00106)	-0.0112*** (0.000448)	-0.0178*** (0.00375)
Mark.Cap.	-0.803*** (0.0237)	-0.752*** (0.00624)	-0.671*** (0.00620)
Leverage	0.0745*** (0.0172)	0.0497*** (0.00726)	0.0333*** (0.00854)
Market Risk Amendment	-20.52* (12.06)	4.101 (10.53)	20.61* (10.95)
Basel 2	-13.80*** (4.711)	-5.467 (3.598)	-1.539 (1.761)
Basel 3	-24.86*** (6.388)	-13.79** (5.918)	12.76 (10.24)
Constant	-61.38*** (13.02)	-49.68*** (11.89)	-43.31*** (11.94)
Year effects	yes	yes	yes
Observations	16,746	16,746	16,746

Clustered standard errors in parentheses

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1

<sup>a</sup> This table reports the results from the .25, .50 and .75 quantile regressions of SRISK. We regress the SRISK measure on the CISS indicator of systemic stress, the market Beta, the internal model dummy from January 1996, the Basel II dummy from June 2006 and the Basel III dummy for September 2008. The Beta is estimated from a GJR-DCC Garch model between the bank stock returns and the MSCI Europe index. We control for year effects, total assets, non-performing assets, other investment in equity and fixed income securities, market capitalization and leverage. The standard errors are clustered for banks (Parente et al. 2016).

We conclude this section looking more carefully at the long-run relationship between SRISK and market capitalization, that could hide a Too-Big-To-Fail story. We perform a VECM analysis on weekly averages with 4 lags, testing for AIC information criteria, LR and test of no-autocorrelation in the residuals. Johansen tests suggest a cointegration equation exists in each subperiod. We perform the VECM on weekly average SRISK, market capitalization, Z-score and EU policy rate. Table 3 reports on the top panel the short-run parameters and reveals that SRISK is a strongly autocorrelated series, significantly affected by short-run movements in Z-score and market capitalization. We do not see endogeneity between SRISK and Z-score, however we do evince endogeneity between market capitalization and systemic risk in the first sub-period. The middle panel reports the cointegrating vector parameters, and we see that both Z-score and the policy rate significantly affect SRISK in a long-run relationship. In the next section, we will also introduce this dynamic behaviour into our specifications.

Table 3: VECM (4 week-lags)

VARIABLES	1987-1996				1996-2006			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	D_SRISKww	D_zscoreww	D_mvwww	D_PolicyRateww	D_SRISKww	D_zscoreww	D_mvwww	D_PolicyRateww
LD.SRISKww	-0.216*** (0.0691)	0.000156 (0.000273)	-0.195*** (0.0531)	5.33e-07 (6.54e-07)	0.205*** (0.0463)	-0.000243** (0.000106)	-0.270*** (0.0453)	-1.46e-07 (1.74e-07)
L2D.SRISKww	-0.0395 (0.0703)	0.000224 (0.000278)	-0.0266 (0.0540)	-4.55e-07 (6.66e-07)	0.0451 (0.0485)	-9.01e-05 (0.000111)	-0.0128 (0.0474)	3.44e-08 (1.82e-07)
L3D.SRISKww	-0.0791 (0.0681)	-0.000261 (0.000269)	-0.141*** (0.0523)	-2.44e-07 (6.45e-07)	-0.0490 (0.0470)	8.06e-05 (0.000107)	-0.0433 (0.0460)	1.69e-07 (1.77e-07)
LD.zscoreww	50.82** (21.62)	-0.282*** (0.0854)	-13.01 (16.61)	0.000358* (0.000205)	-30.08 (20.02)	-0.300*** (0.0456)	-11.32 (19.58)	-3.01e-05 (7.53e-05)
L2D.zscoreww	40.01* (20.44)	-0.0781 (0.0808)	-25.30 (15.71)	0.000485** (0.000194)	-0.829 (20.64)	-0.117** (0.0471)	-8.363 (20.20)	4.09e-05 (7.76e-05)
L3D.zscoreww	1.198 (16.72)	-0.108 (0.0661)	-20.90 (12.85)	0.000243 (0.000158)	17.29 (19.60)	-0.105** (0.0447)	-33.61* (19.18)	6.17e-06 (7.37e-05)
LD.mvwww	-0.178** (0.0847)	0.000759** (0.000335)	-0.277*** (0.0651)	-1.84e-06** (8.03e-07)	-0.172*** (0.0497)	-0.000216* (0.000113)	-0.147*** (0.0487)	-2.36e-07 (1.87e-07)
L2D.mvwww	-0.231*** (0.0876)	0.000342 (0.000346)	-0.327*** (0.0673)	-4.72e-08 (8.30e-07)	-0.0428 (0.0511)	4.56e-05 (0.000117)	-0.0845* (0.0500)	-4.82e-08 (1.92e-07)
L3D.mvwww	-0.156* (0.0865)	0.000751** (0.000342)	-0.141** (0.0665)	-2.00e-06** (8.20e-07)	0.0260 (0.0501)	2.54e-05 (0.000114)	0.00319 (0.0490)	2.50e-07 (1.88e-07)
LD.PolicyRateww	1.303 (6.763)	-21.49 (26.73)	-4.993 (5.197)	0.0524 (0.0641)	4.410 (11.446)	-13.63 (26.10)	5.984 (11.199)	-0.0117 (0.0430)
L2D.PolicyRateww	6.801 (6.732)	13.34 (26.60)	-3.400 (5.173)	-0.0376 (0.0638)	-33.88 (11.452)	-8.418 (26.11)	5.225 (11.204)	-0.0200 (0.0431)
L3D.PolicyRateww	12.363* (6.712)	17.00 (26.53)	-3.197 (5.158)	-0.0490 (0.0636)	-12.320 (11.392)	-15.42 (25.98)	996.8 (11.146)	-0.0214 (0.0428)
Constant	-0.440 (8.778)	-0.0652* (0.0347)	9.254 (6.745)	-8.51e-05 (8.32e-05)	5.471 (9.237)	-0.0108 (0.0211)	7.521 (9.038)	-4.53e-05 (3.47e-05)
L_CE1	-0.0678*** (0.0244)	-0.000507*** (9.66e-05)	-0.00323 (0.0188)	-3.57e-07 (2.31e-07)	-0.0183*** (0.00710)	2.62e-05 (1.62e-05)	0.0133* (0.00694)	-1.10e-07*** (2.67e-08)
CE1 beta	1	888.6288*** (114.963)	0.28429 (0.220)	-24453.22** (10980.8)	1	-750.126*** (197.740)	-0.25925 (0.203)	60611.31*** (23649.9)
Observations	261	261	261	261	541	541	541	541
r2_1	0.169	0.169	0.169	0.169	0.126	0.126	0.126	0.126
r2_2	0.374	0.374	0.374	0.374	0.131	0.131	0.131	0.131
r2_3	0.237	0.237	0.237	0.237	0.0860	0.0860	0.0860	0.0860
r2_4	0.0801	0.0801	0.0801	0.0801	0.0497	0.0497	0.0497	0.0497

<sup>a</sup> This table reports the a VECM of weekly averages of SRISK, market capitalization, Z-score and country policy rates, with 4 lags. We choose the number of lags optimally given the AIC, HQIC and SBIC information criteria and assuring no autocorrelation is left in the residuals. We regress the VECM before 1996 (columns 1-4), and separately between 1996 and 2006 (columns 5-8). Johansen tests of cointegration assure that we have one cointegrating equations in the two periods respectively.

#### 4.4 SRISK and Internal Models of Credit Risk based on Microdata

Overall, so far our results suggest that the January 1996 amendment on market risk had ambiguous effects in reducing systemic risk. While the shift towards internal models apparently was successful in containing systemic risk in the lower quantile of the SRISK distribution, supposedly larger banks in the upper SRISK-quartile tended to benefit less, or even exploit market risk internal models to effectively increase their SRISK positions. Clearly, the aggregate risk enhancing effect dominated the intentional gains on the smaller and less risky banks.

We investigate this self-regulatory tool, looking now at credit risk internal models. One of the pillars of Basel II is the option to widen the scope for internal models also to cover credit risks. While at this stage we do not have sufficiently many (micro) data on the implementation or approval of internal models for market risk, we obtained this micro information about approval and/or adoption of internal credit rating models for a subsample of 100 European banks.<sup>20</sup>

Accordingly, we investigate the relation between SRISK and Basel regulation on the basis of the available bank-level data on the implementation of internal credit risk models. We recall that the variable IRBA takes the value 0 before 2006, the value 1 for the standardized approach, 2 for Foundation-IRBA, 3 for mixed approaches, and 4 for the Advanced-IRBA for credit risk.

Table 4 reports the results of weekly panel mean regressions of SRISK on, alternatively, Beta, CAPM cost of equity, and Delta CoVaR. Banks that implement internal models are larger in size, therefore we include the IRBA dummies with and without interaction with the market value of the bank. It seems that internal models do exert a significant and positive effect on exposure to systemic risk. Especially the mixed approaches seem to contribute increasing the systemic exposure of banks both in mean and in interaction. The choice of using internal models to estimate credit risk contribute to the systemic exposure especially for the largest banks. The market risk amendment dummy is now strongly positively impacting on the systemic exposure.

These results are robust with respect to different models, as different measures of systematic risk or contribution to systemic risk. This evidence stands in stark contrast to the original goals of the Basel Committee in strengthening the safety and soundness of banks.

In the quantile regressions, we address non linearities without an exogenous allocation of banks into the different risk buckets. We focus on the specification with the market Beta, and Table 5 reports the results with the IRBA dummies interacted with market capitalization. The coefficients are estimated via "unconditional" quantile regressions, therefore they can easily interpreted as usual as unconditional effects. We confirm the previous results observing a strong significant impact of internal credit risk models. Very interestingly, the use of mixed or advanced internal model particularly increase the exposure to systemic risk of the institutions in the largest quantile of SRISK. The results are very strong when we look at the intercept effects, while the interacted terms show that it is not only a size-story, but internal models could mitigate the negative effect of size on the exposure to systemic risk of the most systemic banks.

In sum, we find no evidence that the introduction of internal models introduced in 2006 did succeed to increase bank resiliency. The disgression given to the regulated banks apparently, while in compliance with statutory regulation, did not stop banks from engaging in (sophisticated) risk taking activities. The next section aims at showing the robustness of these results in a counterfactual analysis.

So far we confirm the empirical results of Behn and Haselmann (2014) about the unintended consequences of internal models for credit risk for German banks and extend them to European banks

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<sup>20</sup>We are in the process of extending this dataset to all banks. The SNL dataset is relatively small for European banks and needs to be complemented by search for hand-collected implementation dates for the institutions not reported in SNL. We gratefully acknowledge the support of the Bundesbank and the Bank of England in providing the data for all banks in their respective countries. The current subsample is already large enough to allow for meaningful analyses. Behn and Haselmann (2014) analyse an even smaller subsample of German banks.

Table 4: Weekly Panel Regressions of SRISK (residuals, k=0.08%)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Beta	-186.0*** (48.19)	-186.2*** (48.94)	-118.1*** (30.22)						
Cost of Equity				-135.0** (57.05)	-130.0** (56.39)	-125.8** (60.43)			
Delta_CoVaR							26,436*** (5,795)	26,255*** (5,639)	15,950*** (3,902)
CISS	-2.253 (33.06)	-11.65 (31.68)	-9.949 (31.41)	-88.42** (39.54)	-95.69** (38.92)	-83.70** (36.74)	-248.3*** (87.22)	-258.6*** (88.86)	-151.7** (60.14)
Z-Score	-1.136 (1.007)	-1.118 (0.962)	-0.0187 (0.0326)	-0.846 (0.865)	-0.838 (0.775)	0.0288 (0.0257)	0.340 (1.088)	0.278 (0.924)	0.0931* (0.0532)
Market return	-13,267*** (3,448)	-13,239*** (3,439)	-13,309*** (3,460)	-13,308*** (3,481)	-13,284*** (3,474)	-13,296*** (3,478)	-12,576*** (3,246)	-12,552*** (3,238)	-12,910*** (3,335)
Policy rate	822.1** (405.5)	976.2** (427.9)	1,170*** (373.4)	1,765*** (486.2)	1,915*** (508.9)	1,656*** (428.7)	2,482*** (796.4)	2,673*** (815.7)	1,740*** (488.3)
Market-to-Book	0.127*** (0.0111)	0.138*** (0.0118)	0.132*** (0.0109)	0.144*** (0.0101)	0.155*** (0.0124)	0.136*** (0.0109)	0.131*** (0.0178)	0.142*** (0.0190)	0.112*** (0.0239)
MV	0.00248 (0.00192)	0.00560*** (0.00203)	0.00505*** (0.00137)	0.00392* (0.00208)	0.00705*** (0.00202)	0.00384*** (0.00116)	0.00639** (0.00260)	0.00953*** (0.00233)	0.00197 (0.00124)
1.IRBA#MV	-0.00129 (0.00488)		0.00162 (0.00430)	-0.00105 (0.00333)		-0.000532 (0.00299)	0.000397 (0.00279)		-0.00498** (0.00205)
2.IRBA#MV	0.00931** (0.00452)		0.00976*** (0.00333)	0.00993** (0.00403)		0.00918*** (0.00338)	0.0113*** (0.00352)		0.00872** (0.00351)
3.IRBA#MV	0.00460** (0.00201)		0.00185 (0.00201)	0.00463** (0.00230)		0.00300 (0.00183)	0.00452* (0.00262)		0.00494*** (0.00155)
4.IRBA#MV	0.00575 (0.00492)		0.00510 (0.00502)	0.00540 (0.00462)		0.00448 (0.00440)	0.00522 (0.00549)		0.00385 (0.00365)
1.IRBA	-14.33 (14.82)	-8.897 (13.22)	-15.23 (14.25)	18.55* (9.786)	24.11** (10.40)	17.62* (10.32)	51.27** (19.82)	60.04*** (20.95)	47.93** (18.32)
2.IRBA	-34.42 (64.84)	133.2** (63.56)	-51.28* (26.77)	-36.08 (60.90)	143.0** (68.46)	-6.346 (25.04)	-26.97 (51.40)	177.0** (80.92)	41.35 (36.78)
3.IRBA	55.27 (34.27)	130.3** (50.08)	71.72** (31.64)	0.738 (26.30)	75.99* (41.90)	26.48 (21.35)	-64.64* (38.57)	9.044 (44.88)	-44.70 (27.19)
4.IRBA	-44.73 (67.85)	47.54 (61.54)	4.758 (38.76)	-26.97 (47.86)	58.47 (50.03)	8.108 (25.08)	-10.92 (95.32)	70.29 (58.21)	-6.639 (37.81)
Market Amend.	57.47** (26.43)	35.95 (24.20)	22.09 (15.89)	42.52* (24.57)	20.48 (21.46)	50.80*** (16.75)	-8.392 (27.30)	-28.24 (26.32)	64.66*** (21.07)
BEL			-50.84** (21.91)			-65.89*** (19.53)			-69.09* (37.77)
CHE			-30.32 (18.79)			9.175 (12.32)			62.95 (46.50)
CYP			-7.357 (17.08)			8.138 (12.23)			40.87 (47.92)
DEU			10.44 (16.09)			17.68 (12.38)			47.48 (37.90)
ESP			-49.22 (31.98)			-54.15* (32.06)			-88.31* (52.58)
FRA			37.02 (28.75)			41.62*** (14.09)			52.68 (36.58)
GBR			-38.00 (34.54)			-34.73 (32.55)			-18.50 (47.30)
GRC			37.25** (18.75)			7.222 (10.47)			47.09 (37.47)
IRL			-50.39 (37.10)			-59.21* (34.64)			-7.019 (44.88)
ITA			17.01 (17.79)			5.262 (11.87)			-7.699 (38.70)
NLD			-40.42** (18.16)			7.994 (10.69)			31.91 (35.51)
Constant	5.667 (32.52)	-7.302 (32.68)	-45.73* (26.46)	-117.1*** (26.71)	-130.0*** (29.03)	-142.0*** (32.18)	-280.0*** (60.96)	-294.5*** (63.35)	-261.7*** (67.49)
Observations	55,586	55,586	55,586	55,582	55,582	55,582	55,586	55,586	55,586
R-squared	0.014	0.014	0.016	0.012	0.011	0.014	0.019	0.018	0.019
Number of gvkey	95	95		95	95		95	95	

Standard errors in parentheses

\*\*\* p&lt;0.01, \*\* p&lt;0.05, \* p&lt;0.1

<sup>a</sup> This table reports the results from panel regressions with firm fixed effects of the residuals from a first-step AR(2) regression of weekly median SRISK. We regress alternatively, market Beta, CAPM cost of equity, and Delta CoVaR. We include the bank-level IRBA dummies (categories 1 to 4) with and without interaction with the market capitalization of the bank. We also include as regressors the internal model dummy (from January 1996), lagged CISS systemic stress, Z-score, market capitalization, MSCI index and EU country policy rate. The standard errors are clustered for banks.

Table 5: Weekly Unconditional Quantile Regressions of SRISK residuals (k=0.08%)

SRISK	(1) Q.25	(2) Q.50	(3) Q.75	(4) Q.25	(5) Q.50	(6) Q.75
L.SRISK	0.00622*** (0.00235)	0.00619*** (0.00235)	0.169*** (0.0462)	0.0103*** (0.00181)	0.0110*** (0.00202)	0.271*** (0.0411)
L2.SRISK	-0.00632*** (0.00164)	-0.00698*** (0.00164)	-0.161*** (0.0358)	-0.00361** (0.00180)	-0.00351* (0.00202)	-0.0889** (0.0410)
L.Beta	77.72** (34.58)	56.18 (49.12)	3,292*** (687.2)	-5.214 (5.392)	170.5*** (5.194)	4,372*** (93.97)
Zscore	-14.78*** (3.973)	-19.10*** (3.789)	-110.6** (53.17)	-0.0197 (0.0232)	0.340*** (0.0437)	-3.659*** (0.244)
L.CISS	406.4*** (68.94)	339.8*** (68.11)	4,035*** (1,200)	310.2*** (13.98)	254.5*** (14.71)	2,535*** (211.9)
L.market return	-711.1*** (235.0)	-572.9** (224.0)	-9,580*** (2,947)	-649.5** (299.6)	-612.4** (309.7)	-11,634*** (4,350)
L.policy rate	-6,574*** (1,661)	-8,558*** (1,691)	-52,525** (20,667)	-4,451*** (284.0)	-3,803*** (279.1)	-8,343** (4,178)
Market-to-Book	-0.465*** (0.0570)	-0.375*** (0.0977)	-6.981*** (0.431)	-0.551*** (0.0220)	-0.263*** (0.0205)	-3.178*** (0.188)
Market Value	-0.00901** (0.00366)	-0.00973*** (0.00326)	-0.0813* (0.0434)	0.000749** (0.000294)	0.0116*** (0.000332)	0.262*** (0.00578)
1.IRBA-Standardized	22.38 (60.09)	95.56 (62.30)	468.1 (580.5)	176.4*** (8.409)	97.08*** (6.919)	-453.5*** (81.82)
2.IRBA-Foundation	13.92 (155.3)	32.16 (160.3)	860.7 (1,175)	-121.5*** (16.33)	293.5*** (14.50)	-2,519*** (181.2)
3.IRBA-Mixed	227.9* (126.5)	250.5** (122.5)	7,136*** (2,037)	219.6*** (9.790)	440.0*** (9.982)	9,082*** (188.2)
4.IRBA-Advanced	441.4* (235.7)	403.6* (206.3)	589.2 (1,741)	375.5*** (10.81)	879.9*** (12.82)	766.7*** (164.0)
intermodel	-453.9*** (104.1)	-432.3*** (102.5)	-184.4 (2,041)	-521.6*** (16.49)	-657.5*** (16.63)	-4,561*** (273.7)
1.MV#IRBA-Standardized	0.00258 (0.00750)	-0.00136 (0.00758)	-0.0191 (0.135)	-0.0138*** (0.00131)	0.0207*** (0.00117)	0.487*** (0.0219)
2.MV#IRBA-Foundation	0.00688 (0.00582)	0.00582 (0.00569)	0.0817 (0.0734)	0.00532*** (0.000689)	-0.00576*** (0.000704)	0.0444*** (0.0150)
3.MV#IRBA-Mixed	0.00385 (0.00560)	0.00405 (0.00546)	0.0432 (0.0892)	-0.00448*** (0.000361)	-0.0155*** (0.000392)	-0.272*** (0.00714)
4.MV#IRBA-Advanced	-0.0117 (0.00832)	-0.0105 (0.00784)	0.0814 (0.136)	-0.0180*** (0.000744)	-0.0319*** (0.000788)	-0.189*** (0.0140)
Firm Fixed Effects	yes	yes	yes	no	no	no
Country Fixed Effects	no	no	no	yes	yes	yes
BEL				-58.89*** (12.95)	97.53*** (16.40)	3,351*** (382.3)
CHE				-541.4*** (13.53)	-212.9*** (16.07)	-4,143*** (154.2)
CYP				-126.3*** (12.21)	229.9*** (17.29)	-4,715*** (160.9)
DEU				-299.4*** (7.291)	-434.9*** (11.43)	-3,131*** (157.0)
ESP				-379.4*** (12.10)	-206.8*** (14.19)	-3,564*** (239.1)
FRA				0.879 (7.818)	194.7*** (12.15)	3,088*** (214.3)
GBR				-380.6*** (11.13)	-352.0*** (12.94)	-4,473*** (198.9)
GRC				-415.1*** (12.46)	-149.5*** (14.95)	-3,122*** (235.8)
IRL				-424.4*** (15.33)	-194.8*** (16.60)	-4,866*** (225.3)
ITA				-195.1*** (9.059)	-59.60*** (12.85)	-3,916*** (190.8)
NLD				83.65*** (8.049)	533.7*** (13.25)	-7,861*** (221.9)
Constant	750.7*** (171.3)	1,011*** (160.9)	3,866 (2,529)	650.5*** (27.75)	633.6*** (28.81)	4,253*** (430.1)
Observations	56,571	56,571	56,571	56,571	56,571	56,571
R-squared	0.165	0.192	0.261	0.189	0.418	0.572

<sup>a</sup> This table reports the results from the .25, .50 and .75 unconditional quantile regressions of weekly SRISK (Firpo, Fortin and Lemieux, 2009). We include the bank-level IRBA dummies (categories 1 to 4) with and without interaction with the market capitalization of the bank, the internal model dummy (from January 1996). We control for firm effects (1 to 3) or country effects (3 to 6), CISS systemic stress, market capitalization, market investment opportunities proxied by the MSCI equity index and short-term interest rate proxy the country policy rates. The standard errors are clustered for banks (Parente et al. 2016).

and financial institutions. Moreover, we find that the risk enhancing effect of internal models for credit risk are increasing in the systematic importance of banks; in larger and more systemic banks internal models contribute more strongly to an increase in SRISK of European banks and across risk classes. Based on our results the concerns raised about Basel II by Danielson et al. (2001) seem more than justified. By neglecting the endogeneity of systemic risk, Basel II regulation did not succeed to reduce systemic risk ironically precisely in those sectors that turned out to become the most vulnerable ones.

The implementation of Basel II in July 2006 has contributed to moderate the build-up of systemic risk. However, the moderating effect is less striking for precisely the major contributors to systemic risk. In this regard, the speculation of Hakenes and Schnabel (2011) is not supported by the data. Based on theoretical considerations Hakenes and Schnabel argue that the IRB-approach of Basel II induced smaller and medium-sized banks to take larger risks in order to compete effectively with larger banks employing the IRB-approach. We find that their basic assumption that IRB contributes positively to larger banks is not supported by the data.<sup>21</sup>

In summary, the intended consequences of the Basel regulation were achieved only for the safer banks, but ironically they were missed for the riskier banks. Obviously, banks' strategic incentives were not properly understood and the substitutability between capital rules and state guarantees was seriously underestimated throughout the various levels of the Basel process of capital regulation. Consequently, it was especially the systemic European banks that were ill prepared to deal with the subprime crisis in 2007 and even more in the subsequent European sovereign crisis.

#### 4.5 Counterfactual Analysis and Diff-in-Diff Results

By analysing the whole distribution of banks, we go a long way towards causal identification. The documented increase in risk exposure is concentrated on the upper quintiles of the risk distribution of banks, which happen to be the largest and internationally active banks. So the build-up of systemic risk is strongly correlated with the use of internal models, especially for credit risk. Smaller banks are less likely to invest in the costly process of setting up internal credit risk models, and, hence, as we document, build up less exposure to systemic risk. This observation excludes many alternative hypotheses about the increase in risk exposure, such as monetary policy, financial innovation and the role of derivatives. While all these developments undoubtedly influence the strategic choice of business models, our data suggest that it is particularly the large and internationally active banks that exploit those opportunities if at the same time they invest in setting up internal risk models.

In order to assess the contribution of the Basel process to the build-up of risk exposure, we provide a simple counter-factual analysis by asking the question of how would the evolution of risk exposure have occurred if internal models had not existed. In order to conduct this counter-factual experiment, we estimate the behavioural parameters in periods where internal models were not available and then trace the (hypothetical) evolution of systemic risk for the observed realizations of the risk factors in later periods. Thus we can identify the effect of changes in behaviour induced by the

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<sup>21</sup>Even if the competitive effect of Hakenes and Schnabel (2011) is relevant at all, our evidence suggests that the direct (negative) implications for banks' risk management are dominant. However, our findings about the effects of internal models suggest that the assumption of an increase in resiliency of the largest banks due to the use of risk-models is not supported by the data. In this regard, also Colliard (2015) has investigated theoretically the impact of internal models on the risk-taking behaviour of banks.

introduction of internal models.<sup>22</sup>

For conducting the counter-factual experiment we consider two sub-periods: i) the period before any internal models were available (prior to 1996) and ii) the period prior to the implementation of internal credit risk models (1996-2006).

Figure 15 shows the historical evolution of SRISK compared to the estimated forecasted SRISK in case of no changes in the regulatory environment. We present trajectories both for total exposure and average exposure, since the number of banks in our dataset is not constant over time. The results of the mean panel regression suggest that the internal models for market risk did indeed reduce the risk exposure around the turn of the millennium, suggesting that the market risk amendment was in fact helpful in improving risk management for European banks on average.<sup>23</sup> These observations are in line with the original intentions of the Basel Committee even though they appear quantitatively small.

Most strikingly, however, our results suggest that in the run-up of Great Financial Crisis internal models contributed largely to the lack of resiliency of European banks. Our simulations suggest that internal models contributed largely to amplify the capital short-fall in 2008-9 by a factor of two. In fact both, internal models for market risk and for credit risk did contribute to a massive amplification of exposure to systemic risk in the European banking system. After 2014, we observe a significant reduction in aggregate systemic risk, but well above the levels of the Great Financial Crisis. It remains worrisome though that despite improved supervision the capital shortfall remains at the levels of the Great Financial Crisis of an average of about 5 billions of Euros per bank. There is no indication of a normalization of the capitalization of European banks to pre-crisis levels.

How do the reported developments in the banking sector compare to other financial intermediaries, notably the insurance sector but also to real estate funds? Figures 16-17 provide the historical and estimated evolution of SRISK for the various financial sub-sectors: banks, insurance companies and real estates.

As an immediate result, aggregate exposure to systemic risk in banking dominates the other sectors. However, on a per firm basis, the average exposure of the insurance sector is rising and after the Great Financial Crisis even exceeding the average exposure in the banking sector. Moreover, there is strong evidence that capital regulation in the banking sector spills over into the insurance sector, while no spill-overs into the real estate sector can be detected. Nevertheless, the dominant counter-factual effects of capital regulation are clearly identified within the banking sector. This holds both for total as well as average exposure.

Finally, despite a lot of regulatory action and despite the creation of the supervisory infrastructure of Banking Union, there is no tendency for exposure to systemic risk to decline back to pre-crisis levels relative to the Great Financial Crisis.

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<sup>22</sup>This approach follows Fuess et al. (2016).

<sup>23</sup>This is in line with the short-lived reduction in SRISK after the implementation of the internal market risk models in 1996 (see Figure 8).



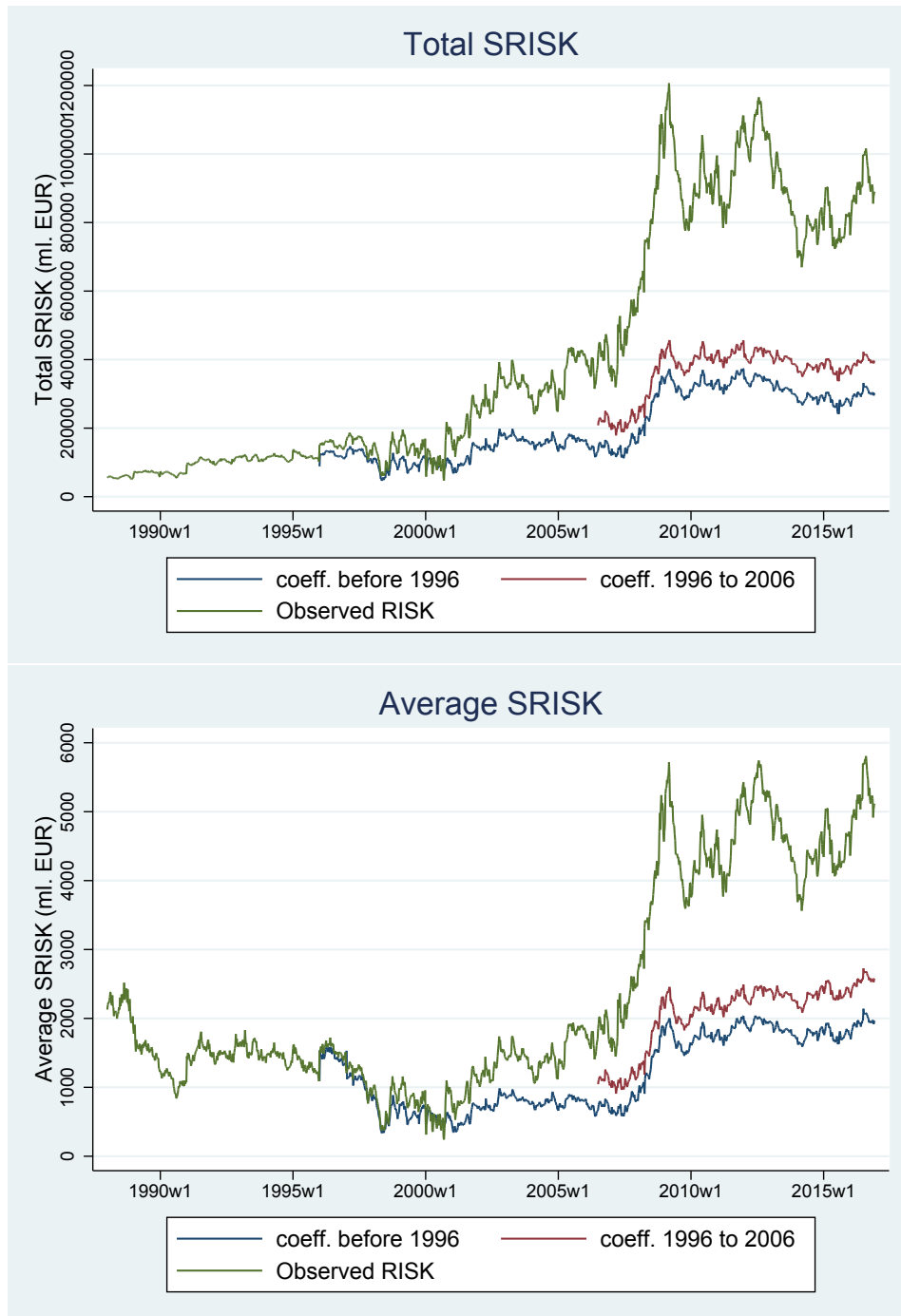


Figure 15: Evolution of historical total and average SRISK vs. counterfactual SRISK. The Figure presents the evolution of the historical average SRISK compared with the estimated forecasted SRISK in case of no Market Risk Amendment (blue line) and no Basel II accord (red line). We estimate SRISK using the dynamic two-stage model as Equation 11 in two sub-periods: i) before any internal models were available (prior to 1996) and ii) prior to the implementation of internal credit risk models (1996-2006).

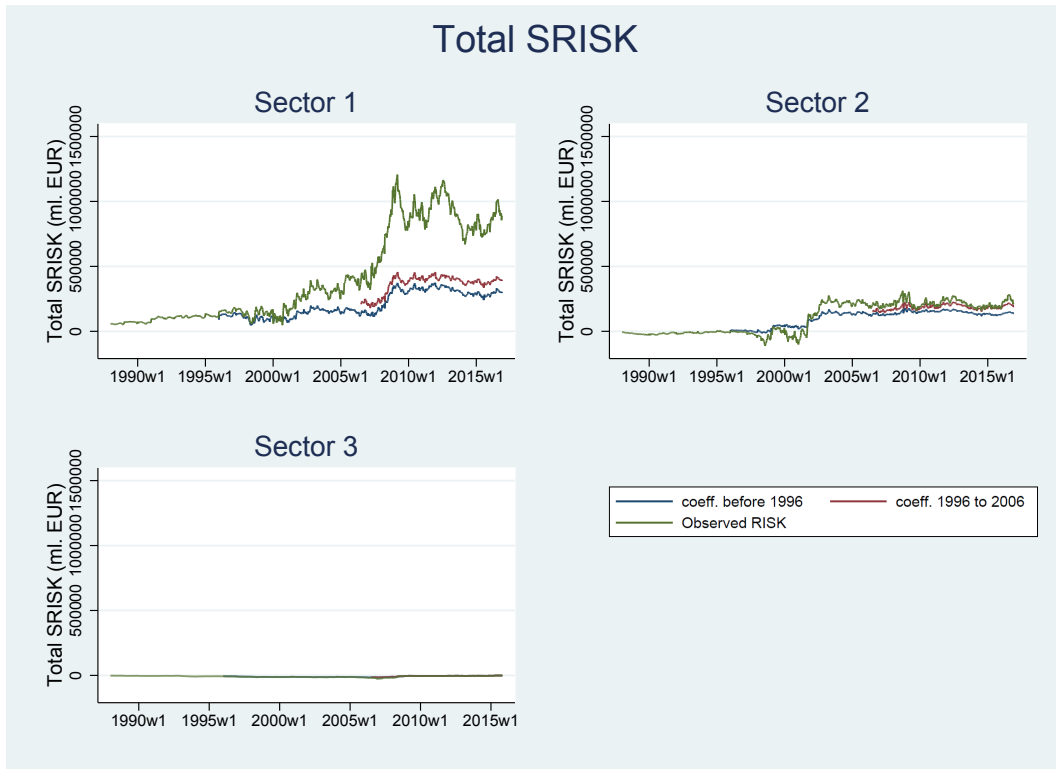


Figure 16: Evolution of historical total sectoral SRISK vs. counterfactual total sectoral SRISK. The Figure presents the evolution of the historical total SRISK compared with the estimated forecasted SRISK in case of no Market Risk Amendment (blue line) and no Basel II accord (red line). We average SRISK according to the financial sector: 1. banks and diversified institutions, 2. insurance companies, 3. real estates. We estimate SRISK using the dynamic two-stage model as Equation 11 in two sub-periods: i) before any internal models were available (prior to 1996) and ii) prior to the implementation of internal credit risk models (1996-2006).

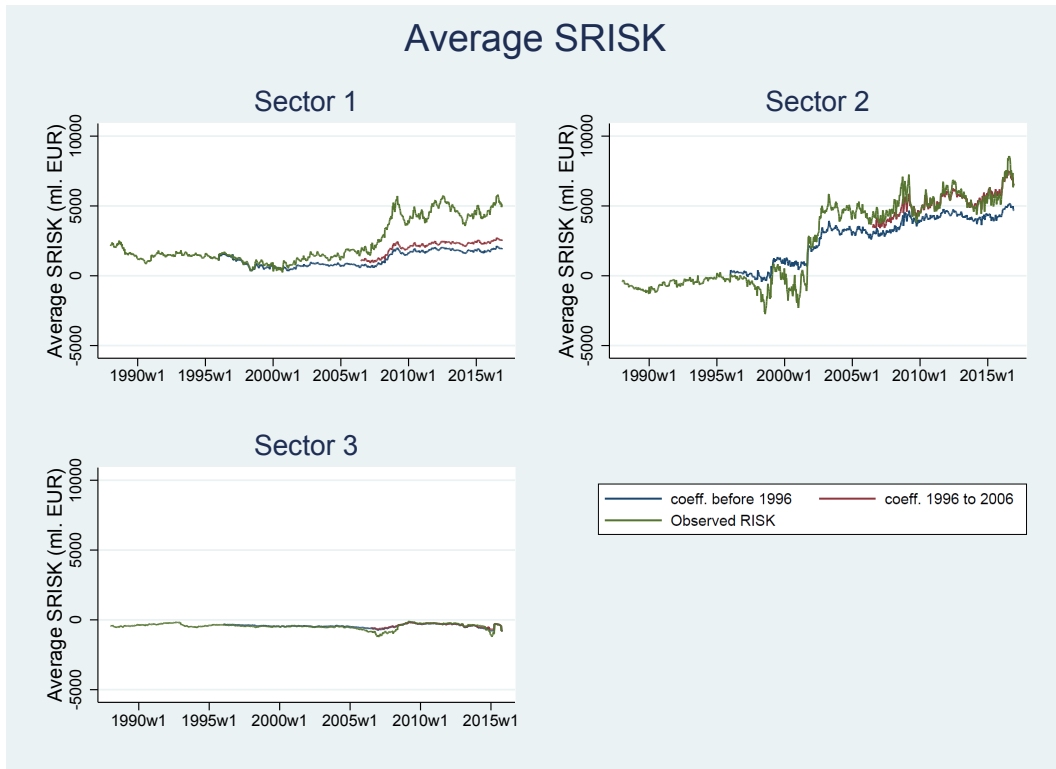


Figure 17: Evolution of historical average sectorial SRISK vs. counterfactual average sectorial SRISK. The Figure presents the evolution of the historical average SRISK compared with the estimated forecasted SRISK in case of no Market Risk Amendment (blue line) and no Basel II accord (red line). We average SRISK according to the financial sector: 1. banks and diversified institutions, 2. insurance companies, 3. real estates. We estimate SRISK using the dynamic two-stage model as Equation 11 in two sub-periods: i) before any internal models were available (prior to 1996) and ii) prior to the implementation of internal credit risk models (1996-2006).

We conclude with a difference-in-differences analysis and report the results on the mean regression. The treated group comprises banks that did implement internal credit risk models after the regulatory option is made available by Basel II in June 2006. We first use propensity score matching (PSM) to weight observations such that the treated group reflects the distribution of covariates in the pre-Basel period. We report both the results for the probit regression for the PSM and the difference-in-differences estimation in Table 6. We choose the bank-characteristics we have used throughout the analysis to identify control banks by lagged SRISK, market beta, Zscore, market-to-book and market value. We see that all these characteristics importantly affect of the choice of implementing internal models, in line with the probit regression results in frame A.

Table 6: PSM and Diff-in-Diff

<b>A. Probit regression</b>				
IRBA01	Coef.	z	p-value	
L.SRISK	-0.0000328***	-20.93	0.000	
L.Beta	0.7261689***	30.59	0.000	
Zscore	-0.0045197***	-4.93	0.000	
Market-to-Book	-0.1658926***	-25.95	0.000	
Market Value	0.0000749***	42.02	0.000	
Constant	-0.4001053***	-15.69	0.000	
Pseudo R2 = 0.2101				
<b>B. Difference-in-differences estimation</b>				
Outcome var.	SRISK	t	p-value	
Baseline:				
Control	5892.937			
Treated	5275.271			
Diff (T-C)	-617.666	-0.22	0.83	
Follow-up:				
Control	1684.41			
Treated	2.00E+04			
Diff (T-C)	1.80E+04	3.3	0.002***	
Diff-in-Diff	1.90E+04	3.73	0.000***	

<sup>a</sup> This table reports the results from the Propensity Score Matching and the following difference-in-differences analysis on banks with internal credit risk models (Advanced or mixed approaches) versus comparable banks without, before and after the regulatory change in 2006. Propensity Score is estimated via a probit regression, where the probability of implementing IRBA is explained by lagged SRISK, market beta, Zscore, Market-to-Book, and market capitalization. We report robust standard errors, clustered per firm. \*\*\*0.01; \*\*0.05; \*0.10.

The PSM provides the weight for the weighted diff-in-diffs regression. The results strongly support our hypothesis that exposure to systemic risk is largely driven by the use of internal credit risk models. While there are no significant differences in SRISK between treatment and control groups prior to the implementation of Basel II standards, we find strong and significant differences in risk exposure after their introduction in 2006. In the follow-up period, we see that institutions that have chosen to implement credit risk models as either advanced or mixed approaches, have increased more than tenfold their exposure to systemic risk in the post-Basel II period compared to the peer group.

## 4.6 Robustness

Our results require a number of important modelling judgements that are open up for discussion. While the specific estimations will depend on these choices, the qualitative results remains remarkably robust. We only provide some of the most immediate request that might come to mind. One suggestion does concern the definition of the prudential capital ratio. While we chose  $k=8\%$  in order to maintain comparability with other studies, we also run regression for the settings  $k=5.5\%$  and  $k=3\%$ . The latter may even open up another interpretation of our risk measure in terms of the leverage ratio under Basel III to which it corresponds. We cannot detect any qualitative modification to the previous discussion, concerning the impact of the Basel dummies and of the credit risk internal models approaches..

While time fixed effects do generate robust and consistent estimations, we also report regressions with country fixed effects instead. For this purpose, we estimate SRISK on the same previous set of regressors via use of OLS or random effects. Overall our results are confirmed with country fixed effects. Systemic risk particularly builds up for the larger more systemic banks.

Concerning monetary policy, again we find weakly destabilizing effect of low interest rates. The estimates are not that strong but again low interest rates reduce resiliency.

Moreover, without over-emphasizing, we find interesting country specific effects, reflecting differences in the supervisory attitudes.<sup>24</sup> Under the mentioned caveat, we find evidence that most countries take a more effective role in reducing SRISK with the notable and significant exception of France. The country effects are significantly lower for Germany and Spain; they are also lower for Italy but not significantly.

## 4.7 Delta CoVaR

We perform the same analysis above using the Delta CoVaR measure from Adrian and Brunnermeier. Following their approach, we use the dollar value of the systemic risk measure, defined as  $\Delta^{\$}CoVaR_{it}(\alpha) = \Delta CoVaR_{it}(\alpha) * size_{it}$ .

As we measure it, Delta CoVaR is the market VaR conditional on a bank being in distress, therefore it can be viewed as measuring an institution's contribution to systemic risk. It measures the contagion deriving from a bank being in distress and, hence, the likelihood of the banking system getting infected by such bank.

Performing a weekly panel regression again we find mixed evidence for the role of the various policy dummies of the Basel regulatory framework (Table 7). While Basel II essentially did succeed to reduce risk of contagion, the implications of the internal models are rather mixed. It turns out that the foundation approach did meet expectations and did contribute to reducing contagion risk especially for the larger banks. However, the mixed approach, which allows bank to apply internal models strategically, tends to contribute positively, and hence enhance, contagion risk. This finding parallels the results on SRISK showing that the strategic use of internal models especially for larger banks contributed to an increase in systemic risk, both for exposure to systemic risk as well as for

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<sup>24</sup>Given data availability and the fact that Banking Union only exists for about 2 years by now - Oct. 2016 - we have not - yet - controlled for a potential structural break in November 2014.

Table 7: Weekly Panel Regressions of  $\Delta^S$  CoVaR

$\Delta^S$ CoVaR	(1)	(2)	(3)	(4)	(5)	(6)
Beta	60.15*** (14.82)	62.07*** (15.48)				
Cost of Equity			-138.5*** (33.32)	-140.3*** (33.82)		
SRISK					0.00246*** (0.000527)	0.00270*** (0.000508)
Market to Book	-0.000139 (0.00385)	-0.00294 (0.00375)	-0.00522 (0.00376)	-0.00803* (0.00406)	-0.00640 (0.00467)	-0.00624 (0.00473)
CISS	99.00*** (24.83)	100.00*** (25.49)	45.62*** (13.62)	45.88*** (13.81)	84.53*** (22.60)	81.66*** (21.51)
MSCI Europe return	-62.79 (55.62)	-51.23 (54.36)	151.9** (65.48)	166.1** (65.48)	-13.66 (55.91)	-24.93 (52.69)
EU Policy rate	-438.4*** (147.3)	-414.9*** (136.7)	-842.4*** (244.6)	-828.7*** (236.5)	-498.3*** (171.0)	-498.7*** (168.6)
Market Value	0.00847*** (0.000467)	0.00758*** (0.000381)	0.00846*** (0.000467)	0.00763*** (0.000437)	0.00842*** (0.000483)	0.00872*** (0.000581)
1.IRBA Standardized #MV		-0.00118 (0.00154)		-0.000989 (0.00204)		-0.00173 (0.00142)
2.IRBA Foundation #MV		-0.00196*** (0.000622)		-0.00197** (0.000760)		-0.00310*** (0.000529)
3.IRBA Mixed #MV		0.00214*** (0.000394)		0.00200*** (0.000384)		-0.000162 (0.000652)
4.IRBA Advanced #MV		0.000498 (0.000367)		0.000283 (0.000512)		-0.00320 (0.00228)
1.IRBA Standardized	-11.96 (11.45)	-9.722 (12.56)	-17.70 (11.44)	-15.90 (13.02)	-6.486 (10.10)	4.923 (9.746)
2.IRBA Foundation	-21.74 (14.06)	-3.749 (12.83)	-24.13* (14.15)	-6.156 (12.87)	-16.53 (14.23)	15.54 (11.00)
3.IRBA Mixed	20.31 (14.94)	-13.13 (11.55)	33.21** (15.80)	2.374 (11.96)	5.878 (14.07)	11.03 (11.60)
4.IRBA Advanced	1.226 (14.50)	-4.177 (14.83)	-1.976 (15.00)	-5.034 (14.62)	-32.74 (22.10)	5.873 (11.86)
intermodel	-32.19 (22.89)	-23.44 (20.21)	-56.32** (26.62)	-48.45** (24.26)	-54.57** (27.43)	-63.91** (25.56)
basel2	-10.32*** (3.049)	-10.02*** (2.985)	-4.507 (2.736)	-4.112 (2.605)	-8.969*** (2.709)	-8.754*** (2.715)
basel3	15.89 (10.54)	15.81 (10.48)	6.622 (9.102)	6.318 (8.874)	10.40 (9.875)	8.407 (9.824)
Year effects	yes*** (p-value)	yes*** (0.0018)	yes*** (0.0008)	yes*** (0.0004)	yes*** (0.0024)	yes*** (0.0013)
Constant	-6.680 (11.09)	-12.46 (10.89)	62.77*** (16.50)	58.78*** (15.04)	36.99*** (13.31)	36.24*** (12.69)
Observations	11,778	11,778	11,778	11,778	11,778	11,778
R-squared	0.538	0.551	0.530	0.542	0.565	0.573
Number of gvkey	100	100	100	100	100	100
rho	0.176	0.176	0.176	0.204	0.114	0.117
r2.b	0.919	0.933	0.958	0.966	0.952	0.951
r2.o	0.769	0.777	0.774	0.776	0.796	0.799
r2.a	0.537	0.550	0.529	0.541	0.563	0.571

Clustered standard errors in parentheses

\*\*\* p&lt;0.01, \*\* p&lt;0.05, \* p&lt;0.1

<sup>a</sup> This table reports the results from panel regressions with fixed effects of weekly median Delta CoVaR in dollar terms. We regress on, alternatively, market Beta, CAPM cost of equity, and SRISK. We include the bank-level IRBA dummies (categories 1 to 4) with and without interaction with the market capitalization of the bank. We also include as regressors the internal model dummy (from January 1996), the Basel II dummy (June 2006 to September 2008) and the Basel III dummy (from September 2008). We control for year effects, CISS systemic stress, market capitalization, MSCI index, and EU country policy rate. The standard errors are clustered for banks.

contributing to systemic risk. Internal models for credit risk were used by larger banks in a way that contributed to enhance systemic risk, and thus reduce stability and soundness for the overall banking system.

Interestingly, the policy rate has strictly negative affects on the contribution to systemic risk, suggesting that a low-growth environment could be an incentive for risk-taking behavior at the level of individual banks.

## 5 Policy Role for Market-Based Risk Measures

An attractive feature of market based risk measures is the fact that they control for market feedback. Regulatory institutions and supervisors, however, typically focus on information about individual institutions and, hence, idiosyncratic risk such as possibly embodied in book values. Also the Basel capital regulation focuses on book rather than market values. This regulatory approach, while facilitating the analysis of single institutions by separating them from market developments, is not helpful in a system context, since the very foundations of systemic risk are tied to the notion of market feedback. Bank runs do occur because of depositors' (self-fulfilling) fears about other depositors running. Contagion effects occur, whenever insolvencies of single institutions cause knock-on insolvencies of connected, but otherwise healthy, financial institutions. Accordingly, discrepancies between book and market values may contain important systemic information to which supervisors (and regulators) should not cast a blind eye.

In order to illustrate the informational content of the market based capital shortfall measure SRISK, we provide a brief discussion of two systemic European banks that entered into different trajectories during the Great Financial Crisis, Deutsche Bank and Union Bank of Switzerland (UBS) (Figures 18). While UBS had to be rescued by the tax payers in 2007, Deutsche Bank succeeded to (narrowly) escape the need of government support in 2007-8. In the respective SRISK trajectories we identify similar pre-crisis developments. Both banks had accumulated a pre-crisis shortfall of about 60 bill. Euro according to our crisis definition. During the crisis the measure shot up to about 160 bill. Euro in the case of Deutsche Bank, while in the case of UBS the tax payer intervened and the measure only increased to about 100 bill. Euro. In 2010 the shortfall measures declined in both cases but remained considerable above pre-crisis level until the European Sovereign Crisis hit, increasing the short fall again for both banks. But even after 2013 in the case of Deutsche Bank the capital shortfall basically remained at level of 2009, considerably above the pre-crisis level of 2007. In contrast UBS succeeded in reducing capital shortfall to pre-crisis levels of 2007 and even below.

The recent troubles of Deutsche Bank after the leakage of hefty penalties in the United States in September 2016 are clear evidence that capital shortfall is strongly correlated with lack of investor confidence and a high degree of stock market volatility, essentially due to worries about the bank's resilience. Quite differently, UBS seems to stay out of trouble quite comfortably despite the realizations of operational risk also on their side.

European supervisors tend to take the view that markets may be over-reacting to bad news causing market-to-book values to be excessively depressed. They seem to be essentially satisfied by what they consider serious attempts of Deutsche Bank - and other systemic banks - to rebuild

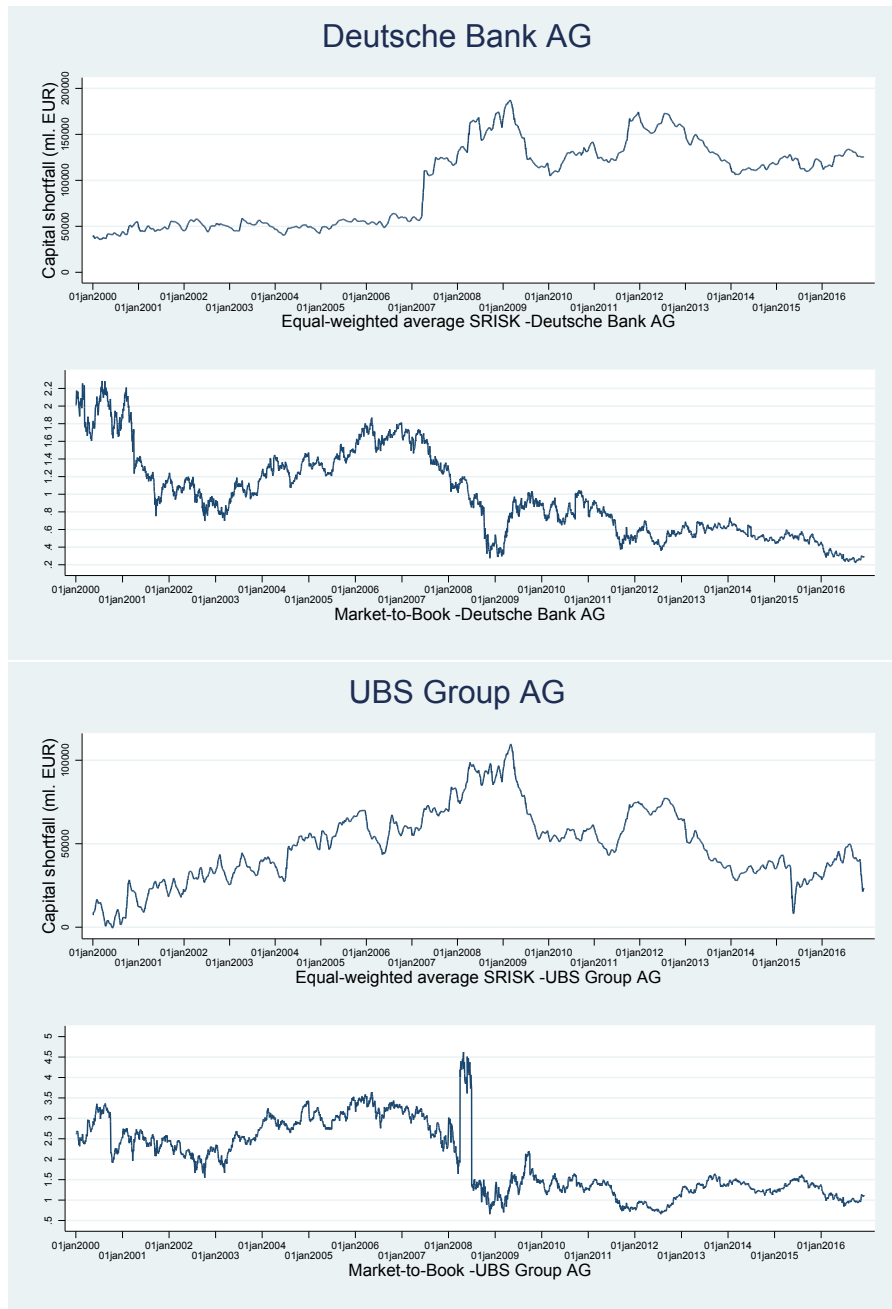


Figure 18: Cross-sector variation of systematic risk. The Figure presents the evolution of the SRISK and Market-to-Book of Deutsche Bank and UBS.

*book values* of regulatory tier-1 capital <sup>25</sup>. Also the ECB tends to be more concerned to harmonize supervisory procedures for smaller banks than to recapitalize the ailing systemic banks in Europe

<sup>25</sup>See e.g. Carney, 2016, Dombrovskis, 2016 and Nouy, 2016



(see Gehrig et al. 2016). The case of UBS is an interesting case study, since i) Switzerland is over-complying with Basel III standards, and ii) UBS is over-complying with Swiss standards. And in fact, market-to-book recovered for UBS to essentially normal values, while in the case of Deutsche Bank, market-to-book remains on a long run decline well below .5.<sup>26</sup> The case of UBS demonstrates that it is possible to rebuild market confidence and, thus, market valued capital, if the recapitalization is done seriously enough. Obviously, it is very costly to undo the massive stock repurchases in the run-up to the Great Financial Crisis, but rebuilding confidence requires serious and similarly massive commitment. Market values are important indicators of market confidence and trust, and, hence, relevant information also for supervisors.

## 6 Concluding Comments

While the Basel process of capital regulation was designed to increase the stability and safety of the global banking system, we provide evidence that it did not fully achieve this goal in its first three decades of operation for European banks. From the perspective of systemic risk measures, the Basel process has been effective for smaller banks in our sample. But even there it did not significantly reduce systemic exposures or contagion risk. For the largest quantiles of banks, internal models might have provided strong incentives to carve out equity and, thus, reduce in-house resiliency. The evidence demonstrates that those incentives were exploited and the resiliency of large and especially systemic European banks was greatly impaired at the onset of the Great Financial Crisis. To the extent that all large banks did engage in this activity of reducing their capital buffers, overall bank capital also became scarce, generating systemic concerns for the whole banking sector. But even after 2008, and especially even after the start of the Banking Union in November 2014, most individual - and thus aggregate - SRISK scores did not retreat to levels of the 2007-8 crises or below.<sup>27</sup>

Our empirical evidence is consistent with the view that risk models were chosen strategically (see Behn, Haselmann and Vig (2014) and Colliard (2015), but also Admati and Hellwig (2013)) resulting in an enormous depletion of bank equity. Ironically, these equity carve-outs were one way of increasing return on equity through extensive stock repurchases especially prior to the Great Financial Crisis and at a time when the cost of bank equity was actually low, and strengthening capitalization and resiliency would have been relatively cheap (in historical context).<sup>28</sup>

On the basis of our analysis it is not necessarily that capital rules per se were insufficient; it is rather the possibility to reduce effective capitalization by means of complex risk models under supervisory approval that causes the lack of resiliency. Our findings accord well with Miles et al. (2012). They seem to contradict Jackson (2015) in the sense that simple models, even at sub-optimal levels in terms of efficiency, may be more suitable to limit risks and, hence, safeguard resiliency.

It is not evident that these outcomes should be viewed as unintended consequences of the Basel process of capital regulation. Rather public warnings about such outcomes had dutifully and rig-

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<sup>26</sup>In September market-to-book for Deutsche Bank even fell as low as .10.

<sup>27</sup>This observation is consistent with attempts of ECB researchers (Homar, Kick, Salleo, 2016) trying to empirically validate the ECB policy of focusing on particular on the European ECB and EBA stress scenarios rather than focusing on individual and aggregate capital shortfall for the Euro area as suggested for example by Acharya, Engle and Pierret (2014).

<sup>28</sup>Baron and Xiong (2014) provide a behavioural explanation based on over-optimism.

ously been voiced by leading academic researchers. Notably Danielson et al. (2001) raise serious concerns that the neglect of endogeneity of systemic risk could turn into an unintended build-up of major systemic risk within the Basel II approach. However, the political economy of the Basel process might have succumbed to industry interests in reducing the bite of the standard approach, and introducing options to determine regulatory capital with the help of internal models.

The current debate on Basel IV reflects a new focus on this front, implicitly showing that regulators have become aware of the need to limiting the potential misuse of internal models. The Basel Committee's Consultative Document on credit risk models (Basel Committee on Banking Supervision, 2016) proposes to remove this auto-regulatory option for exposures that do not allow for sufficiently reliable estimates, such as low-default exposures.

We also suggest that, by concentrating on formal fulfilment of regulatory rules based on book values, regulators missed a pro-social role in interpreting (negative) market feedback. Relying on rules based on book values only completely neglects social feedback and market expectations. However, trust and confidence are key in the banking industry, but they are notoriously difficult to measure and observe. Hence, market based risk measures are one simple step towards taking into account market reactions, trust and confidence, and hence systemic market feedback. This is potentially useful information and supervisors should be challenged more when disregarding market information.<sup>29</sup> After all, supervisors attain an important role to correct potential misbehavior only in market economies. This argument assumes the existence of a sufficiently high degree of trust in the operation of markets after all. If this trust is not given, why not economize on bureaucracy and centralize the whole banking system?

Our analysis also uncovers disconcerting effects of monetary policy on banks' contribution to systemic risk. We can motivate this result considering that a low-growth environment might create incentives for excessive risk-taking, and, therefore, increase both, contagion risk and contribution to systemic instability of the sector. This effect can be observed throughout the whole distribution of banks. Accordingly, under the current regulatory framework, Quantitative Easing, through its effect on interest rates, might tend to contribute to undermining the stability and soundness of the European banking system.<sup>30</sup>

There are even wider implications of the Basel process of capital regulation beyond the banking industry on the whole financial sector (Gehrig, Iannino, 2016). For example, the build-up of systematic risk in the insurance sector, while not as dramatic as in the banking sector, also significantly moves upwards with a structural break around 1996. Possibly these developments also exhibit unintended consequences across markets and industries: long-term lending is increasingly given up by banks<sup>31</sup> and taken over by the insurance sector. Hence, a final evaluation of the welfare consequences of the Basel process of capital regulation requires an analysis whole financial sector in order to not only account for market feedback, both in the regulated as well as the unregulated segments, but also for substitution effects and their implications on complementary activities. We

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<sup>29</sup>This argument is not saying that there is no mispricing in markets. However, under normal conditions mispricing should be a short term problem. In the long run markets should converge to fair valuations. For example, a market-to-book anomaly may occur for short periods; but when it persists for years or decades, the underlying sources of the anomaly may be important to remedy.

<sup>30</sup>This finding suggests that Quantitative Easing would require complementary supervisory instruments to control adverse risk-taking incentives. In the case of Europe such complementary control was not effective for the period of our study.

<sup>31</sup>On the shortening of banks' planing horizon see also Boot and Ratnovski (2016).

leave this for future research.

We leave for future research also the interaction between capital regulation and Banking Union. It is too early for a final judgement of Banking Union on the most systemic banks. However, at this stage we cannot detect any decline in the systemic risk scores for the banks under direct ECB supervision. Certainly, their SRISK remain well above the 2008 levels still in 2016.

Similarly, the implications of Brexit on financial stability are too early to assess. While for a final judgement the process of Brexit still needs to be properly defined, early market reactions do suggest a slight increase in systemic risk only. This seems to reflect an increase in risk premia due to heightened uncertainty. Due to the fact that systemic risk of the most systemic European banks has not been checked effectively up to date under Banking Union, Brexit implications for the stability of the European banking system may still be formidable and serious.

We end with the observation that the build-up in systemic risk in the financial sector entails considerable tail risk for the macro economy, which has been identified as one likely channel for secular stagnation (e.g. Kozlowksi, Veldkamp, Venkatesvaran, 2015).<sup>32</sup> To the extent that one subscribes to this argument, it is true that the Basel process has contributed to permanently enhancing tail risk. Thus, under this view, real effects of the resulting equity carves out in European and global banking systems can be seen contributors to the decline in long-term investment growth. The "missing recovery" after the Great Depression (Summers, 2016, Teulings Baldwin, 2014), unfortunately, correlates strongly with high levels of systemic risk, particularly for the largest, and, presumably, most efficient financial institutions worldwide.

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<sup>32</sup>Kozlowski et al. (2015) argue that rare event realizations of tail risk have changed long term beliefs and expectations. Analysing credit spreads Füss et al. (2016) find similar evidence of changing risk perceptions in the U.S. corporate debt market.

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